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1 **Increased Frequency of Summer Extreme Heat Waves over Texas Area Tied to**
2 **the Amplification of Pacific Zonal SST Gradient**

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ABSTRACT

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Summer extreme heat waves (EHWs) over Texas area and their trend are investigated using observations and atmospheric general circulation model (AGCM) outputs. There is a positive linear trend in Texas EHW days for the period of 1979-2015. While the interannual variability of the Texas EHWs is linked to ENSO conditions, the upward trend in Texas EHWs is found to be significantly associated with the tropical Pacific Zonal SST Gradient (PZSSTG). The amplification of PZSSTG leads to both enhanced convection in the western Pacific and suppressed convection in the central-eastern Pacific (i.e., La Niña-like pattern), both of which can induce anomalous anticyclones over the Texas area through two distinct planetary wave trains in the antecedent spring. As a result, anomalously sinking motions and divergent water vapor flux appear over the Texas area, which reduce precipitation and increase downward solar radiation, leading to a dry and hot soil that favors the occurrence of Texas summer EHWs. In addition, all AGCMs using observed SSTs as boundary conditions were able to simulate the observed decreasing trend in Texas summer precipitation and the observed increasing trend in Texas summer surface air temperature. The observed relationships between winter PZSSTG and the following spring-summer Texas precipitation/temperature were also reproduced by these models, where the intensified PZSSTG tended to reduce the Texas precipitation while increasing the surface air temperature.

41 **1. Introduction**

42 Texas and its surrounding areas have experienced numerous extreme heat waves (EHWs)
43 over the past decades, which have produced large impacts on human health, agricultural
44 productions, and natural ecosystem. For example, an EHW swept the central and southern
45 United States in 1980, causing 107 heat-related deaths in Texas (Karl and Quayle 1981;
46 Greenberg et al. 1983). In 1998, a more localized EHW struck Texas and Oklahoma that led
47 to an estimated loss of 6 billion US dollars primarily due to decreased agricultural production
48 (Chenault and Parsons 1998; Hong and Kalnay 2000). In 2011, an unprecedented EHW
49 occurred over the Texas area, with an average temperature almost 3°C above the 1981–2010
50 mean for June through August (Nielsen-Gammon 2012; Hoerling et al. 2013), which
51 substantially increased the emergency department visits for heat-related illnesses (Zhang et al.
52 2015). Among these cases, the 2011 EHW was the most notable in both intensity and duration,
53 and was accompanied by a record-breaking burned area in the southern and southwestern
54 Texas due to wildfires (Williams et al. 2014). Smith et al. (2013) indicated that the EHW
55 frequency, in terms of EHW days per year, over southern North America, experienced an
56 upward trend based on the various EHW definitions. More intense and longer lasting EHWs
57 over the Texas area have raised concerns of the potential impact of greenhouse warming on
58 the increasing frequency of Texas EHWs.

59 The potential mechanisms associated with EHWs include the precipitation – evaporation
60 – temperature feedback (e.g., Fischer et al. 2007; Lorenz et al. 2010; Mueller and
61 Seneviratne 2012) and persistent anticyclones or blocking highs (e.g., Dole et al. 2011;

62 Trenberth and Fasullo 2012; Screen and Simmonds 2014). Both of these can be triggered or
63 forced by the remote sea surface temperature anomalies (SSTAs) or internal atmospheric
64 dynamics. For example, Hong and Kalnay (2000) found that the Pacific SSTAs could
65 establish large-scale conditions for Texas drought during the antecedent spring through
66 atmospheric teleconnection. The spring soil moisture anomalies subsequently play an
67 important role in maintaining the drought and triggering the summer EHWs by a positive
68 feedback associated with lower evaporation/precipitation. On the other hand, Lyon and Dole
69 (1995) analyzed the large-scale circulations associated with the Texas EHW in 1980, and
70 found that this particular event was primarily forced by a stationary wave propagating
71 southeastward from an apparent source region south of the Aleutians. Using an atmospheric
72 general circulation model (AGCM), Teng et al. (2013, 2016) indicated that the EHWs in the
73 United States tend to be preceded by a pattern of anomalous atmospheric planetary waves
74 with a wavenumber of 5 by 15–20 days. Petoukhov et al. (2013) further noted that the 2011
75 Texas EHWs were significantly connected with the planetary waves with zonal
76 wavenumbers 6, 7, or 8 that are trapped within the mid-latitude waveguide. Screen and
77 Simmonds (2014) also proposed that the amplification of quasi-stationary waves with
78 zonal wavenumbers 3–8 preferentially increases the probability of EHWs in North America.

79 Other studies have compared the roles of greenhouse warming and oceanic forcing in
80 affecting the drought/heat waves in Texas area. Using an AGCM, Rupp et al. (2012, 2015)
81 investigated the influence of anthropogenic greenhouse warming on the Texas EHW in 2011
82 and concluded that the likelihood of exceeding a given unusually high summer temperature in

83 the Texas region was about 10 times greater with 2011 anthropogenic emissions compared to
84 pre-industrial forcing. Furthermore, Rupp et al. (2013, 2017) assessed the influences of
85 greenhouse gases and ocean's role for the 2012 central United States drought and found that
86 the SSTAs, rather than the anthropogenic forcing, were more likely to increase the occurrence
87 of the 2012 drought/heat. Wang et al. (2014) compared the roles of SST forcing in the 2011
88 and 2012 drought and heat events in the United States using the NASA Goddard Earth
89 Observing System version 5 (GEOS5) AGCM, and found that the winter/spring responses
90 over the United States to the Pacific SSTAs were remarkably similar for these two years
91 despite substantial differences in the tropical Pacific SST, implying that the SSTAs outside
92 the central and eastern Pacific might also play some roles.

93 It is well known that precipitation deficits in southern and southwestern North America
94 are linked to the tropical Pacific SSTA, notably to the cold state of the eastern Pacific, which
95 usually leads to anticyclonic anomalies over these regions that favor high pressures and dry
96 conditions (e.g., Schubert et al., 2004; Seager and Ting 2017). However, there is no
97 conclusive evidence showing whether La Niña activities have been enhanced or damped in
98 recent decades due to the relatively small samples of the ENSO events (Collins et al. 2010).
99 Thus ENSO alone may not be sufficient to explain the increasing Texas EHWs.

100 Hoerling and Kumar (2003) linked the drought/heat in the United States to the cooling in
101 the eastern tropical Pacific and the warming in the western Pacific. The warmth of the
102 Indo-Pacific oceans has been unprecedented in recent decades, accompanied by an enhanced
103 Pacific Zonal SST Gradient (PZSSTG) (L'Heureux et al. 2013; McGregor et al. 2014). A

104 strengthened PZSSTG, on the one hand, favors the cold state maintenance in the central –
105 eastern Pacific, which can induce robust anticyclones over southern North America and the
106 occurrence of EHWs. On the other hand, the intensified PZSSTG can also enhance
107 convective activities in the western tropical Pacific, which may also affect the North
108 American EHWs through teleconnection. Thus, it is of interest to determine the possible
109 impacts of ENSO and the PZSSTG on the increasing trend in Texas EHWs and to explore the
110 underlying mechanisms.

111 In the current study, the trend and year-to-year variability in Texas EHWs are
112 investigated. We also compare the different physical processes associated with ENSO and
113 PZSSTG, focusing mainly on the tropical western Pacific. The rest of the paper is organized
114 as follows. In section 2, we describe the data sets and analysis methods. In section 3, we
115 discuss the overall features of EHWs in North America. In section 4, we explore the drivers
116 and associated mechanisms. The AGCM simulated results are discussed in section 5,
117 followed by a summary in section 6.

118 **2. Data and Method**

119 *a. Observations and model outputs*

120 For comparison purposes, we use two sets of data for SST, precipitation, and maximum
121 2-m temperature (Mx2t). The monthly mean SST data sets are obtained from the National
122 Oceanic and Atmospheric Administration (NOAA) extended reconstructed version 4
123 (NOAA-ERA v4; Huang et al. 2015) and the Hadley Centre (HadISST, Rayner et al. 2003),
124 with horizontal resolutions of $2^{\circ}\times 2^{\circ}$ and $1^{\circ}\times 1^{\circ}$, respectively. The land precipitation data sets

125 are acquired from the Global Precipitation Climatology Centre (GPCC; Schneider et al. 2011)
126 and the European Centre for Medium-Range Weather Forecasts (ECMWF) ERA-Interim
127 (Dee et al. 2011), both having a horizontal resolution of $0.5^{\circ} \times 0.5^{\circ}$. The daily Mx2t data sets
128 are used to define EHWs, which are obtained from the ECMWF ERA-Interim (Dee et al.
129 2011) and the NCEP-DOE Reanalysis 2 (provided by the NOAA/OAR/ESRL PSD, Boulder,
130 Colorado, USA; their Web site: <http://www.esrl.noaa.gov/psd/>), with a resolution of $0.5^{\circ} \times 0.5^{\circ}$
131 and a global T62 Gaussian grid (192×94), respectively. For most figures, the Texas EHWs
132 and their relations with other variables are shown based on the ERA-Interim data.

133 The atmospheric variables, including geopotential height, three-dimensional velocity at
134 17 levels, the 4-layer volumetric soil moisture, and the vertical integrals of eastward and
135 northward moisture fluxes, are used to diagnose the associated large-scale conditions for the
136 variation in EHWs. These data sets, with a horizontal resolution of $2.5^{\circ} \times 2.5^{\circ}$, are also
137 obtained from the ECMWF ERA-Interim (Dee et al. 2011). The Niño3.4 index is defined as
138 the area-averaged SSTAs over the central and eastern Pacific ($170^{\circ}\text{E} - 120^{\circ}\text{W}$, $5^{\circ}\text{S} - 5^{\circ}\text{N}$)
139 based on the NOAA-ERV4 data. The PZSSTG is computed from the NOAA-ERV4 SSTA
140 differences between the tropical western Pacific ($120^{\circ}\text{E} - 150^{\circ}\text{E}$, $10^{\circ}\text{S} - 10^{\circ}\text{N}$) and Niño3.4.
141 The analysis period in this study is 1979–2015.

142 To assess the impacts of ENSO and PZSSTG on Texas precipitation and temperature, we
143 analyze the outputs from several AGCMs provided by the NOAA Drought Task Force (DTF;
144 Schubert et al., 2009). The reason these DTF experiments were chosen here is that the models
145 tend to have good skill at simulating different drought mechanisms, feedbacks, and potential

146 predictability of several high-profile cases (Wood et al. 2015). These cases include the
147 southeastern U.S. drought during 2006/07, the Texas drought of 2011, the central Great Plains
148 drought of 2012, and the western U.S. drought from 1998 to 2002. The model outputs used in
149 this study are from the AGCM experiments forced by observed SST for the period of 1979–
150 2014. The models are ECHAM5, CFSv2, CAM4, GEOS5, and CCM3, whose horizontal
151 resolutions are $0.75^{\circ}\times 0.75^{\circ}$, $1^{\circ}\times 1^{\circ}$, $1.25^{\circ}\times 0.75^{\circ}$, $1.25^{\circ}\times 1^{\circ}$, and $2.5^{\circ}\times 2.5^{\circ}$, respectively. Each
152 AGCM produced 12-20 ensemble members. We calculate the ensemble means for each
153 model before analyzing the results.

154 *b. EHW thresholds and EHW days*

155 The EHWs are defined by a percentile-based threshold method, which is widely used
156 (e.g., Meehl and Tebaldi 2004; Della-Marta et al. 2007; Kuglitsch et al. 2010). For a specific
157 day within the June–August period, its maximum temperature threshold is determined by the
158 95th percentile of Mx2t for a total of 555 days (37 years \times 15 days, the 15 days correspond to
159 7 days on either side of the target date) for the 37-year period from 1979 to 2015. By moving
160 the 15-day window forward or backward, we get the consecutive threshold for every target
161 date. An EHW event is identified by two criteria: (1) there are at least 3 consecutive days that
162 the Mx2t exceeds its 95th percentile threshold, and (2) the average Mx2t during the EHW
163 event must exceed 30°C. Therefore, the EHW days (EHWD) can be acquired by computing
164 the total days of EHWs over a specific period (an example is given in Fig. 2c for the 2011
165 summer in Texas). A large EHWD represents that there are more threshold-breaking hot days
166 or more frequent EHWs, and vice versa.

167 In addition, we also use the singular value decomposition (SVD) method to explore the
168 relationship between Pacific SSTAs and North American EHWD, which allows us to identify
169 their covariability (Bretherton et al. 1992). In sections analyzing the interannual variability in
170 Texas EHWs, the linear trends and low-frequent signals in Texas EHWs are removed by the
171 10-year running mean method. The statistical significances of the composite and correlation
172 results are tested by the two-tailed Student's *t*-test, with a degree of freedom of 35 for a total
173 of 37 years (1979–2015).

174 **3. Characteristics of EHWs over North America**

175 Figure 1a shows the climatological patterns of North American Mx2t and 500-hPa
176 geopotential height averaged for the boreal summer (JJA). We can see that there exist two
177 high temperature centers, which are located over southwestern and south-central North
178 America, where Mx2t exceeds 33°C. The hot regions coincide with the subtropical
179 anticyclone, which stretches from the subtropical North Atlantic to the eastern Pacific. It is
180 well understood that a subtropical high is often accompanied by descending air motion, less
181 precipitation, and clear skies, thus resulting in high daily maximum temperature.

182 Figure 1b presents the total EHWD in North America during JJA for the period 1979–
183 2015. In general, the relatively high value of EHWDs can be found over western,
184 southwestern, and south-central North America, compared to the northern and eastern regions.
185 Such distribution is expected given the North American climate landscape of dry west–wet
186 east as well as the frequently reported droughts over the regions with large EHWDs (e.g., Yin
187 et al. 2014; Rupp et al. 2015). The largest values of EHWD seem to appear over

188 Texas/Oklahoma areas, where there are more than 80 summer days (out of 3404 days) during
189 JJA when the Mx2t exceeds the EHW threshold.

190 The linear trend in EHWD is shown in Fig. 1c. Although strong EHW activities exist
191 over western North America, the trend there is insignificant. In comparison, the largest trend
192 in EHWD is clearly seen over Texas and its nearby areas, which is significant at the 95%
193 confidence level. The EHWD trends shown in Fig. 1c are consistent with many previous
194 studies (e.g., Lau and Nath 2012; Smith et al. 2013; Teng et al. 2016), which reported that the
195 largest trends in EHWD occurred in southeastern North America and the Great Plains.
196 However, these studies did not address why the largest trend in EHWD appeared in the Texas
197 area. Therefore it is our goal here to determine why the Texas areas seem to be the preferred
198 locations of increasing EHWs and what are the underlying mechanisms?

199 To answer these questions, we further investigate the year-to-year variations of Texas
200 EHWs and associated atmospheric and oceanic conditions. Figure 2a shows the area-averaged
201 EHWD over the domain (105°W – 90°W , 28°N – 38°N) based on both 90% and 95% thresholds
202 over the 37 years. While large interannual variation exists in the Texas EHWD, the 10-year
203 running mean (dashed line) seems to indicate a consistent upward trend from 1979 to 2015.
204 The largest EHWD was found in 2011, which has been recognized as one of the most extreme
205 summers (e.g., Zhou et al. 2014; Rupp et al. 2015; Zhang et al. 2015). As a way to remove the
206 trend and focus on interannual variation, we identified 10 years when the values of EHWD
207 exceeded the 10-year running mean (i.e., 1980, 1986, 1989, 1996, 1998, 2000, 2006 2009,
208 2011, and 2012, see supplementary Fig. S1 for each year's EHW Days spatial patterns), as

209 marked in Fig. 2a. These 10 years are selected to make the composite analysis next.

210 Figure 2b shows the monthly Niño3.4 index and area-averaged precipitation for Texas
211 and surrounding area (105°W – 90°W , 28°N – 38°N , blue box in Fig. 1c), from January 1979 to
212 December 2015. The 10 severe EHW summers are indicated by the black triangles in Fig. 2b,
213 which correspond well with the central and eastern Pacific SST cooling during the preceding
214 winter and spring. Further examinations reveal that most winters (8 out of 10) preceding the
215 Texas summer EHWs are featured by anomalously cold SSTAs in the central and eastern
216 Pacific, except for the 1979/80 and 1997/98 winters (See supplementary Fig. S2). As seen
217 from Fig. 2b, large precipitation deficits are also found preceding each occurrence of Texas
218 EHWs. In fact, Texas precipitation and Niño3.4 are well correlated for all months from
219 January 1979 to December 2015, with a correlation coefficient of 0.47, suggesting that the
220 central and eastern Pacific cooling can lead to significant rainfall deficit in spring and summer,
221 which further lead to drought in the Texas region in the summer (Seager and Hoerling 2014).

222 The 2011 summer is seen as the case with the most extreme EHWD for the entire study
223 period, which follows a strong La Niña event with reduced spring and summer Texas
224 precipitation (Fig. 2b). We show in Fig. 2c a selected site (98°W , 34°N), which is a grid cell
225 near Wichita Falls, Texas, to illustrate the evolution of this extreme case based on
226 ERA-Interim data. As can be seen, the 95th threshold temperature is generally lower in the
227 beginning and the end of the summer, and peaks in mid-summer. Most of the dates above the
228 95% thresholds during this summer were between mid-June and late-August, when there were
229 consecutive days with maximum Mx2t exceeding 40°C (104°F).

230 To determine statistically whether the increase in Texas daily Mx2t was simply due to
231 the background warming (thus a shift of the mean Mx2t) or it might involve additional
232 processes, Fig. 3a shows the probability density function (PDF) of Texas daily Mx2t for the
233 10 extreme summers (red), the rest of the 27 summers (grey), the first half of the period from
234 1980-97 (blue) and the second half of the period from 1998 to 2015 (cyan). The PDFs for the
235 two different periods (blue vs. cyan) show a clear shift due to changes in the mean, from 32°C
236 in the first half to 35°C in the second half, which could reflect a global warming contribution
237 or the Pacific/Atlantic multi-decadal modulation (McCabe et al. 2004) to the Texas EHWs.
238 The PDF for the 10 extreme summers, however, show a more dramatic increase in probability
239 of the high daily Mx2t compared to all three other PDFs in Fig. 3a. Compared to the normal
240 years, i.e., the rest of the 27 years, the possibility of occurrence of Mx2t above 35°C is much
241 higher in these 10 extreme summers. In particular, the probability of occurrence of Mx2t
242 between 36°–39°C is substantially higher in extreme summers. Thus, the increased EHWs in
243 the Texas area could be affected not only by the shift in mean temperature, but could be also
244 influenced by additional processes involved in causing these extremes. In Fig. 3b, we
245 illustrate that the seasonal mean Mx2t and the EHWD averaged over the Texas area are
246 highly correlated with a correlation coefficient of 0.87. This linear relationship is more
247 evident for the 10 EHWs summers. Therefore, in cases when daily Mx2t is not available, it is
248 an alternative method to use the seasonal mean surface temperature to represent EHWs in the
249 Texas area.

250 **4. Drivers and mechanisms associated with Texas EHWs**

251 *a. Relationship between Pacific SSTA and Texas EHW*

252 It is clear from Fig. 2b that there is a strong relation between Texas EHW and SSTA in
253 the tropical Pacific. Figure 4 explores this relationship further by compositing tropical Pacific
254 SSTAs and 200-hPa geopotential height in the preceding winter, spring, and simultaneous
255 summer based on the 10 most EHW events of Texas. We can see that warming and cooling
256 SSTAs appear in the western and central Pacific, respectively, resembling the mature and
257 decaying La Niña patterns from winter to summer. Correspondingly, as shown in the right
258 panels of Fig. 4, stationary Rossby waves are triggered in the upper troposphere, which
259 originate from the central Pacific and propagate poleward. The stationary waves branch out in
260 two directions over the North Pacific, one continuing across the Arctic region and dissipating
261 in the eastern hemisphere, and the other turning southeastward in both winter and spring. By
262 the spring season, an anomalous high pressure is found over the Texas area, which persists
263 into the summer season.

264 The precipitation and vertically-integrated water vapor flux (WVF) composites for the
265 EHW years are shown in Fig. 5, for both the spring and summer seasons. During the
266 antecedent spring, the westward and eastward WVF anomalies prevail over the subtropical
267 and mid-latitude regions, respectively, due to the anomalous atmospheric anticyclones, which
268 results in significant precipitation deficit in southern North America. During summer, the
269 divergent WVF anomalies intensified, and were accompanied by anomalous southward WVF
270 leading to widespread precipitation deficit in Texas and its surrounding regions.

271 The results in Figs. 4 and 5 suggest a plausible mechanism for the Texas EHWs under

272 La Niña conditions. A La Niña event first causes anomalous downstream stationary wave
273 propagation, which induces anticyclonic anomalies over the Texas area in spring and summer.
274 The anomalous anticyclone then reduces cloud cover and precipitation, which leads to drier
275 soils, and thus fosters more EHWs there (Hong and Kalnay 2000). It is, however, unclear if
276 the La Niña SSTAs could have caused the increasing trend in Texas EHWs. Previous studies
277 indicated that the PZSSTG had intensified over the past three decades, which was
278 accompanied by the recent strengthening of the Pacific Walker circulation (e.g., Kosaka and
279 Xie 2013; McGregor et al. 2014). Is the increasing trend in Texas EHWs the result of
280 intensified PZSSTG or purely due to La Niña conditions?

281 To answer this question, Fig. 6 shows the time series of the PZSSTG and Niño3.4, as
282 well as their relations with Texas precipitation. We can see that the PZSSTG presents a
283 significant strengthening trend in both ERSST and HadiSST data sets, while the trend in
284 Niño3.4 is small and insignificant, implying that the western Pacific SST warms faster than
285 that in the central and eastern Pacific in observations. As shown in Figs. 6c and 6e, the
286 intensification of PZSSTG tends to significantly decrease the spring and summer precipitation
287 over the Texas area. The correlation coefficient between antecedent winter PZSSTG and the
288 following spring (summer) Texas precipitation reaches -0.45 (-0.40), exceeding the 99%
289 (95%) confidence level. As a special case, the precipitation anomaly during the 2011 summer
290 is marked in Fig. 6e, which coincided with the largest PZSSTG.

291 In comparison, the scatter plots between Texas precipitation and the Niño3.4 index are
292 shown in the right panels of Fig. 6. The correlation between Texas precipitation and the

293 winter Niño3.4 index decreases from -0.45 to -0.33 (from -0.40 to -0.36) during the boreal
294 spring (summer), though still significant above the 95% confidence level. The 2011 case is
295 marked in Fig. 6f, showing the Texas precipitation deficit corresponded to a moderately large
296 Niño3.4 value, but not the strongest La Niña.

297 Figure 7 further shows the relationships between DJF PZSSTG (Niño3.4) and Texas area
298 soil water and between DJF PZSSTG (Niño3.4) and Mx2t. It should be noted that although
299 significant linkages between these variables are found, the uncertainties in these values are
300 relatively large, especially for the soil water provided by ERA-Interim. Therefore, one should
301 be more cautious in explaining these correlation relationships. The summer Texas soil water
302 shows a significantly negative correlation with the winter PZSSTG, with a correlation
303 coefficient of -0.42, which could be related to the precipitation deficits in previous seasons.
304 Figure 7c shows that the Texas Mx2t tends to increase corresponding to the amplification of
305 DJF PZSSTG, which likely results from the precipitation – soil moisture – temperature
306 feedback (Schär et al. 1999; Fischer et al. 2007). The 2011 summer was extremely hot, which
307 was accompanied by the largest precipitation deficit and dry soil condition. Similar to the
308 Texas precipitation, the correlation coefficient between Niño3.4 and Texas soil water/Mx2t is
309 less significant compared to that with PZSSTG. In fact, the correlation coefficient between
310 DJF negative Niño3.4 (i.e., multiplied by -1) and Texas JJA Mx2t is only 0.29, which is
311 below the 95% confidence level. Thus, although the DJF La Niña may play an important role
312 in the occurrence of Texas summer EHWs, the western Pacific SSTAs can also contribute to
313 some extent, which enhances the PZSSTG and the associated tropical convection.

314 *b. Physical mechanisms associated with Texas EHW*

315 How much can the Pacific SSTAs and associated convections explain the trend and
316 variability of Texas EHWs? To answer this question, we performed Singular Value
317 Decomposition (SVD) analysis between the Pacific outgoing longwave radiation (OLR) in
318 DJF and the North American EHWD in the following summer (JJA) to extract the
319 co-variability between tropical convection and extreme heat waves. The OLR is usually used
320 to measure the convective intensity in the tropics, where a small (larger) OLR value indicates
321 stronger (weaker) convection. As seen from Fig. 8a, the first SVD mode presents negative
322 correlation in the western Pacific (WP), positive correlation in the central Pacific (CP), and
323 negative correlation in the eastern Pacific (EP), implying enhanced convection over the WP
324 and suppressed convection over the CP. Meanwhile, the OLR pattern is correlated with
325 significant warming in the WP and cooling in the CP, implying an intensification of PZSSTG.
326 It is well known that an intensified PZSSTG could drive a stronger Walker circulation in the
327 equatorial Pacific, which favors the WP convection. Thus, the enhanced convection over the
328 WP may be viewed as a response to the increased PZSSTG.

329 Correspondingly, the EHWD over most of the US, particularly in Texas and Oklahoma,
330 increases (Fig. 8b), which accounts for 47.8% of the total covariance between OLR and
331 EHWD. Figure 8c shows the time series of the first SVD mode for OLR and EHWD, both
332 showing an upward trend, with a correlation coefficient of 0.63 between the two variables,
333 indicating a direct link between enhanced convection over WP and the Texas EHWs. To
334 better understand the dynamical linkage between WP convection and Texas EHWs, a WP

335 OLR index is constructed using the area-averaged OLR over the WP domain (100°E–130°E,
336 15°S–20°N).

337 The WP OLR and PZSSTG indices are shown in Fig. 9a. The WP OLR shows a
338 significantly downward trend, implying that the WP convection became stronger in recent
339 decades. In comparison, the PZSSTG shows a significantly increasing trend, which
340 corresponds well with the decrease in WP OLR, with a high correlation coefficient of -0.78.
341 To determine the atmospheric circulation features associated with the WP convection, we
342 show in Fig. 9(b) the regression pattern of 200-hPa geopotential height onto the negative WP
343 OLR index (i.e., multiplied by -1). Corresponding to the enhancement of WP convection,
344 there exist two distinct Rossby wave trains. One originates from the central Pacific, similar to
345 that shown in Fig. 4, which is likely triggered by ENSO SST anomalies. The other originates
346 from the WP and propagates across the North Pacific, contributing to the anomalously high
347 pressure over southern North America during boreal spring. After removing the ENSO signal
348 by regressing out Niño3.4 from the WP OLR index, the Rossby waves originating from the
349 central Pacific disappear almost completely (Fig. 9c), while the WP originated Rossby waves
350 intensified. By comparing Figs. 9b and 9c, one can see clearly that both La Niña and the
351 enhanced convection in the WP contribute to the anticyclones and thus EHWs over the
352 southern United States.

353 The anomalous circulation pattern associated with WP OLR tends to persist to the
354 summer (See supplementary Fig. S3), although the magnitude of the wave train weakens,
355 which results in anomalous high pressure over the Texas area, leading to prolonged dry and

356 hot conditions. The correlation coefficient between the spring (summer) Texas precipitation
357 and the ENSO – removed negative WP OLR index is approximately -0.3 (-0.2), which is
358 significant at the 90% (80%) confidence level. The summer Texas Mx2t also shows a positive
359 relationship (0.28) with the ENSO – removed negative WP OLR index, which confirms the
360 hypothesis that the WP convection may contribute to the occurrence of Texas EHWs through
361 wave propagation, although the relationship is not as strong as the one including ENSO.

362 **5. AGCM experiment results**

363 Five AGCMs from the NOAA Drought Task Force simulations (Schubert et al., 2009)
364 are used in this section to further examine the linkage between Texas climate and Pacific SST
365 conditions. Forced by observed SST, these AGCMs simulated well the climatology of
366 summer precipitation and 2-m temperature (T2m) over North America, with dry west–wet
367 east pattern regardless of their different horizontal resolutions (see supplementary Fig. S4).

368 Figure 10 shows the ensemble means of precipitation and T2m averaged over the Texas
369 area during the summer. All five AGCMs simulated a negative trend in Texas precipitation
370 since 1990s and a positive trend in T2m. Due to the lack of daily model outputs, the EHWs
371 cannot be defined in the same way as before. Instead, we use monthly mean T2m temperature
372 to define extreme hot summers, a reasonable alternative considering the close relationship
373 between seasonal mean Mx2t and EHWD as shown in Fig. 3b. The increasing trend in Texas
374 summer T2m simulated by all models implies more frequent EHWs over Texas area in Fig.
375 10b. The correlation coefficient between the multi-model mean (MMM) T2m and
376 precipitation in summer is as high as -0.9, indicating that the increase in Texas surface air

377 temperature primarily results from the reduction in local precipitation. In comparison, the
378 correlation coefficient between Texas surface air temperature and precipitation from the
379 observation is -0.84, only slightly smaller than that from the model simulations. The small
380 reduction in correlation is due to the larger atmospheric internal variability in observations as
381 compared to the multi-model ensemble mean. The high correlation confirms that lower
382 precipitation could lead to drier soil conditions and reduced evaporation, leading to higher
383 surface temperature and less precipitation. In addition, the decreased precipitation is usually
384 accompanied by fewer clouds and more surface solar radiation, which lead to a higher surface
385 air temperature.

386 Figure 11 shows the regression maps of simulated spring 200-hPa geopotential height
387 onto the antecedent winter PZSSTG. One of the most prominent features is the La
388 Niña-forced wave pattern, showing anomalous low-pressure centers over the tropical Pacific
389 and northern North America, and anomalously high-pressure centers over the North Pacific
390 and southern North America. After linearly removing the ENSO signals, the WP-originated
391 wave trains become more obvious, which propagate across the North Pacific and lead to
392 anomalously high pressure over the Texas area. Note that substantial heterogeneity exists in
393 the intensity and location of the simulated wave trains among different AGCMs, although the
394 anomalous anticyclones over the Texas area were simulated by most AGCMs, implying that
395 both the WP and EP SSTAs exert influences on Texas climate.

396 We further examine the relationships between Texas precipitation/temperature and
397 PZSSTG/Niño3.4 within the AGCM framework. The left panels of Fig. 12 show that the

398 winter PZSSTG has a negative correlation with the MMM precipitation over Texas during the
399 following spring (summer), with a correlation coefficient of -0.82 (-0.73). The correlation
400 coefficients between PZSSTG and Texas precipitation simulated by the AGCMs are much
401 higher than that in observations, as expected from the multi-model mean that reduces
402 atmospheric internal variability and emphasizes the SST-forced signal. Correspondingly, the
403 PZSSTG shows a positive correlation with the MMM T2m over Texas during summer, with a
404 correlation coefficient of 0.77. The AGCM results confirm that the amplification of winter
405 PZSSTG tends to reduce precipitation and increase T2m over the Texas area.

406 The right panels of Fig. 12 show the similar scatter plots between Texas
407 precipitation/temperature and negative Niño3.4 index. While the DJF SSTAs in the central-
408 eastern Pacific also show significant relationships with Texas precipitation and temperature in
409 the following spring and summer, the magnitudes of the correlation coefficients are much
410 reduced compared with those for the PZSSTG. These results confirm again that the WP
411 SSTAs and the associated wave trains contribute to the variability in precipitation and
412 temperature over the Texas area.

413 **6. Summary**

414 This study investigated the trend and year-to-year variability in Texas extreme heat wave
415 events. In climatology, the EHWs mainly occur over western, southwestern, and southern
416 North America, as revealed by previous studies (e.g., Lau and Nath 2012; Smith et al. 2013;
417 Teng et al. 2016). The largest trend of EHWs is found over the Texas area, which is found to
418 be related to the enhanced tropical PZSSTG in the antecedent winter and spring. The

419 enhanced PZSSTG can be a result of La Niña conditions, which lead to the cooling of eastern
420 tropical Pacific and result in anticyclonic circulation anomalies over the Texas areas through
421 planetary wave propagation. However, La Niña alone is found to be insufficient to explain the
422 increasing trend in Texas EHWs, suggesting the possible contributions of the warming
423 western tropical Pacific to the Texas EHWs.

424 The SVD analysis between tropical Pacific OLR and North American extreme heat wave
425 days reveals that the WP convection associated with the amplification of PZSSTG is
426 correlated with the increased frequency of Texas EHWs. In recent decades, the winter
427 PZSSTG has experienced an increasing trend, consistent with the Texas EHWs. The winter
428 PZSSTG also shows a significant correlation with the Texas precipitation in following spring
429 (summer). In comparison, the correlation coefficients between Niño3.4 and the Texas
430 precipitation/temperature in following spring and summer are weaker than that with the
431 antecedent PZSSTG, suggesting that the warm WP SSTAs may contribute to the Texas
432 EHWs.

433 The physical mechanisms linking the occurrences of Texas EHWs and the PZSSTG are
434 found to be associated with two distinct wave trains. One is triggered by La Niña conditions
435 and originates from the central tropical Pacific that propagates northeastward, leading to an
436 anomalous anticyclone over the Texas area. The other originates from the western tropical
437 Pacific and propagates northeastward across the North Pacific, which also contributes to the
438 Texas high pressure anomaly. Although the La Niña related wave trains dominate the Texas
439 climate, the WP originated wave trains add significantly to the relationships between

440 PZSSTG and the Texas climate. The increased PZSSTG favors the intensification of WP
441 convection and the maintenance of stronger Pacific cooling, both of which contribute to the
442 anticyclone anomalies over Texas area. Under the control of a persistent anticyclone, sinking
443 air and clear skies prevail, which suppress the local convection and reduce precipitation.
444 Through precipitation – soil moisture – temperature feedback, higher surface temperatures
445 and the increased occurrences of EHWs are expected.

446 Our study points at the importance of the anomalous WP SSTAs and convection to the
447 Texas spring and summer precipitation as well as the summer heat wave events. Previous
448 studies (e.g., Kosaka and Xie 2013; McGregor et al. 2014) have found an enhanced PZSSTG
449 as a result of greenhouse warming based on the coupled ocean–atmosphere model simulations,
450 which would point to possible future increases in Texas heat wave occurrences and
451 exacerbated drought in the region through atmosphere teleconnection. It should be cautioned
452 that our study does not explicitly address the role of anthropogenic forcing in recent Texas
453 droughts/heats. Rupp et al. (2015) indicated that no simulated increase in the frequency of
454 large precipitation or soil moisture deficits was detected from preindustrial to year 2011
455 conditions. The dynamic mechanism proposed here could also apply to shorter time scales,
456 such as the intraseasonal time scales, when the western tropical Pacific convection due to the
457 Madden–Julian Oscillation (e.g., Barlow and Salstein 2006; Zhou et al. 2012) can trigger
458 similar wave trains and lead to Texas heat waves.

459

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468 for Climate, Environment and Ecosystem.

REFERENCES

469

470 Barlow, M., and D. Salstein, 2006: Summertime influence of the Madden–Julian oscillation

471 on daily rainfall over Mexico and Central America. *Geophys. Res. Lett.*, **33**, L21708.

472 Bretherton, C. S., C. Smith, and J. M. Wallace, 1992: An intercomparison of methods for

473 finding coupled patterns in climate data. *J. Climate*, **5**, 541–560.

474 Chenault, E., and G. Parsons, 1998: Drought worse than 96; cotton crop's one of worst ever.

475 <http://agnews.tamu.edu/stories/AGEC/Aug1998a.htm>, Texas A&M Agricultural News

476 Home Page, College Station, TX, August 19.

477 Collins, M., and Coauthors, 2010: The impact of global warming on the tropical Pacific and

478 El Niño. *Nat. Geosci.*, **3**, 391–397.

479 Dee, D. P., and Coauthors, 2011: The ERA-Interim reanalysis: Configuration and

480 performance of the data assimilation system. *Quart. J. Roy. Meteor. Soc.*, **137**, 553–597.

481 Della-Marta, P. M., M. R. Haylock, J. Luterbacher, and H. Wanner, 2007: Doubled length of

482 western European summer heat waves since 1880. *J. Geophys. Res.*, **112**, D15103.

483 Dole, R., and Coauthors, 2011: Was there a basis for anticipating the 2010 Russian heat wave?

484 *Geophys. Res. Lett.*, **38**, L06702, doi:10.1029/2010GL046582.

485 Fischer, E. M., S. I. Seneviratne, P. L. Vidale, D. Lüthi, and C. Schär, 2007: Soil moisture–

486 atmosphere interaction during the 2003 European summer heat wave. *J. Climate*, **20**,

487 5081–5099.

488 Greenberg, J., J. Bromberg, C. Reed, T. Gustafson, and R. Beauchamp, 1983: The

489 epidemiology of heat-related deaths, Texas–1950, 1970–79, and 1980. *AJPH*, **73**, 805–

490 807.

491 Hoerling, M., and A. Kumar, 2003: The perfect ocean for drought. *Science*, **299**, 691–694.

492 Hoerling, M., and Coauthors, 2013: Anatomy of an extreme event. *J. Climate*, **26**, 2811–2832.

493 Hong, S. Y., and E. Kalnay, 2000: Role of sea surface temperature and soil-moisture feedback
494 in the 1998 Oklahoma–Texas drought. *Nature*, **408**, 842–844.

495 Huang B., V. F. Banzon, E. Freeman, J. Lawrimore, W. Liu, T. C. Peterson, T. M. Smith, P. W.
496 Thorne, S. D. Woodruff, and H.-M. Zhang, 2015: Extended Reconstructed Sea Surface
497 Temperature version 4 (ERSST.v4). Part I: Upgrades and intercomparison. *J. Climate*, **28**,
498 911–930.

499 Karl, T. and R. Quayle, 1981: The 1980 summer heat wave and drought in historical
500 perspective. *Mon. Weath. Rev.*, **109**, 2055–2073.

501 Kosaka, Y., and S.-P. Xie, 2013: Recent global-warming hiatus tied to equatorial Pacific
502 surface cooling. *Nature*, **501**, 403–407.

503 Kuglitsch, F. G., A. Toreti, E. Xoplaki, P. M. Della-Marta, C. S. Zerefos, M. Türkeş, and J.
504 Luterbacher, 2010: Heat wave changes in the eastern Mediterranean since 1960.
505 *Geophys. Res. Lett.*, **37**, L04802.

506 Lau, N.-C., and M. J. Nath, 2012: A model study of heat waves over North America:
507 Meteorological aspects and projections for the twenty-first century. *J. Climate*, **25**,
508 4761–4784.

509 L’Heureux, M. L., S. Lee, and B. Lyon, 2013: Recent multidecadal strengthening of the
510 Walker circulation across the tropical Pacific. *Nat. Climate Change*, **3**, 571–576.

511 Lorenz, R., E. B. Jaeger, and S.I. Seneviratne, 2010: Persistence of heat waves and its link to
512 soil moisture Memory. *Geophys. Res. Let.*, **37**, L09703.

513 Lyon, B., and R. M. Dole, 1995: A diagnostic comparison of the 1980 and 1988 U.S. summer
514 heat wave–droughts. *J. Climate*, **8**, 1658–1676.

515 McCabe, G. J., M. A. Palecki, and J. L. Betancourt, 2004: Pacific and Atlantic Ocean
516 influences on multidecadal drought frequency in the United States. *Proc. Nat. Acad. Sci.*,
517 **101**, 4136–4141.

518 McGregor, S., A. Timmermann, M. F. Stuecker, M. H. England, M. Merrifield, F.-F. Jin, and Y.
519 Chikamoto, 2014: Recent Walker circulation strengthening and Pacific cooling amplified
520 by Atlantic warming. *Nat. Climate Change*, **4**, 888–892.

521 Meehl, G. A. and C. Tebaldi, 2004: More intense, more frequent, and longer lasting heat
522 waves in the 21st century. *Science*, **305**, 994–997.

523 Mueller, B., and S. Seneviratne, 2012: Hot days induced by precipitation deficits at the global
524 scale. *Proc. Natl. Acad. Sci. USA*, **109**, 12398–12403.

525 Nielsen-Gammon, J., 2012: The 2011 Texas drought. *Texas Water J.*, **3**, 59–95.

526 Peterson, T. C., and Coauthors, 2013: Monitoring and understanding changes in heat waves,
527 cold waves, floods, and droughts in the United States: State of knowledge. *Bull. Amer.*
528 *Meteor. Soc.*, **94**, 821–834.

529 Petoukhov, V., S. Rahmstorf, S. Petri, and H. J. Schellnhuber, 2013: Quasiresonant
530 amplification of planetary waves and recent Northern Hemisphere weather extremes.
531 *Proc. Nat. Acad. Sci.*, **110**, 5336–5341.

532 Rayner, N. A., and Coauthors, 2003: Global analyses of sea surface temperature, sea ice, and
533 night marine air temperature since the late nineteenth century. *J. Geophys. Res.*, **108**,
534 4407.

535 Rupp, D. E., P. W. Mote, N. Massey, C. J. Rye, R. Jones, and M. R. Allen, 2012: Did human
536 influence on climate make the 2011 Texas drought more probable? *Bull. Am. Meteorol.*
537 *Soc.*, **93**, 1041–1067.

538 Rupp, D. E., P. W. Mote, N. Massey, F. E. L. Otto, and M. R. Allen, 2013: Human influence
539 on the probability of low precipitation in the Central United States in 2012. *Bull. Am.*
540 *Meteorol. Soc.*, **94**, S2–S6.

541 Rupp, D. E., S. Li, N. Massey, S. N. Sparrow, P. W. Mote, and M. R. Allen, 2015:
542 Anthropogenic influence on the changing likelihood of an exceptionally warm summer
543 in Texas, 2011. *Geophys. Res. Lett.*, **42**, 2392–2400.

544 Rupp, D. E., S. Li., P. W. Mote, N. Massey, S. N. Sparrow, and D. C. H. Wallom, 2017:
545 Influence of the ocean and greenhouse gases on severe drought likelihood in the central
546 US in 2012. *J. Climate*, **30**, 1789–1806.

547 Schär, C., D. Lüthi, and U. Beyerle, 1999: The soil-precipitation feedback: A process study
548 with a regional climate model. *J. Climate*, **12**, 722–741.

549 Schneider, U., and coauthors, 2014: GPCCC’s new land surface precipitation climatology based
550 on quality-controlled in situ data and its role in quantifying the global water cycle. *Theor.*
551 *Appl. Climatol.*, **115**, 15–40.

552 Schubert, S. D., M. J. Suarez, P. J. Pegion, R. D. Koster, and J. T. Bacmeister, 2004: On the

553 cause of the 1930s Dust Bowl. *Science*, 303, 1855–1859,
554 doi:<https://doi.org/10.1126/science.1095048>.

555 Schubert, S., D. Gutzler, and Coauthors, 2009: A US CLIVAR Project to Assess and Compare
556 the Responses of Global Climate Models to Drought-Related SST Forcing Patterns:
557 Overview and Results. *J. Climate*, **22**, 5251–5272.

558 Screen, J. A., and I. Simmonds, 2014: Amplified mid-latitude planetary waves favour
559 particular regional weather extremes. *Nat. Climate Change*, **4**, 704–709.

560 Seager, R., and M. Hoerling, 2014: Atmosphere and ocean origins of North American
561 droughts. *J. Climate*, **27**, 4581–4606.

562 Seager, R., and M. Ting, 2017: Decadal drought variability over North America: mechanisms
563 and predictability. *Cur.r Climate Change Rep.*, **3**, 141–149.

564 Smith, T. T., B. F. Zaitchik, and J. M. Gohlke, 2013: Heat waves in the United States:
565 definitions, patterns and trends. *Climate Change*, **118**, 811–825.

566 Teng, H., G. Branstator, H. Wang, G. Meehl, and W. Washington, 2013: Probability of US
567 heat waves affected by a subseasonal planetary wave pattern. *Nat. Geosci.*, **6**, 1056–
568 1061.

569 Teng, H., G. Branstator, G. A. Meehl, and W. M. Washington, 2016: Projected intensification
570 of subseasonal temperature variability and heat waves in the Great Plains. *Geophys. Res.*
571 *Lett.*, **43**, 2165–2173.

572 Trenberth, K. E., and J. T. Fasullo, 2012: Climate extremes and climate change: The Russian
573 heat wave and other climate extremes of 2010. *J. Geophys. Res.*, **117**, D17103.

574 Wang, H., S. Schubert, R. Koster, Y.-G. Ham, and M. Suarez, 2014: On the role of SST
575 forcing in the 2011 and 2012 extreme U.S. heat and drought: A study in contrasts. *J.*
576 *Hydrometeor*, **15**, 1255–1273.

577 Williams, P. A., and Coauthors, 2013: Temperature as a potent driver of regional forest
578 drought stress and tree mortality. *Nat. Climate Change*, **3**, 292–297.

579 Wood, E. F., S. D. Schubert, A. W. Wood, C. D. Peters-Lidard, K. C. Mo, A. Mariotti, and R.
580 S. Pulwarty, 2015: Prospects for advancing drought understanding, monitoring and
581 prediction. *J. Hydrometeor.*, **16**, 1636–1657.

582 Yin, D., M. L. Roderick, G. Leech, F. Sun, and Y. Huang, 2014: The contribution of reduction
583 in evaporative cooling to higher surface air temperatures during drought. *Geophys. Res.*
584 *Lett.*, **41**, 7891–7897.

585 Zhang, K., T.-H. Chen, and C. E. Begley, 2015: Impact of the 2011 heat wave on mortality
586 and emergency department visits in Houston, Texas. *Environ. Health*, **14**, 11.

587 Zhou, S., M. L’Heureux, S. Weaver, and A. Kumar, 2012: A composite study of the MJO
588 influence on the surface air temperature and precipitation over the continental United
589 States. *Climate Dyn.*, **38**, 1459–1471.

590 Zhou, W., S. Ji, T. H. Chen, Y. Hou, and K. Zhang, 2014: The 2011 heat wave in greater
591 Houston: Effects of land use on temperature. *Environ. Res. Lett.*, **2014**, 135.

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595 **Figure Captions**

596 **Figure 1.** (a) Climatological maximum 2-m temperature (Mx2t) and 500-hPa geopotential
597 height (H500, unit: m). (b) Total summer extreme heat wave (EHW) days (EHWD) over the
598 period of 1979–2015. (c) Linear trends in EHWD. The blue box in (c) outlines the Texas area
599 (105°W – 90°W , 28°N – 38°N). Variables in (a)–(c) are obtained from ERA-I product.

600 **Figure 2.** (a) Light blue curve and grey bar indicate the 90th and 95th percentile EHWDs,
601 respectively. The dashed curve denotes the 10-year running mean of the 90th percentile
602 EHWD. The 10 EHW years (1980, 1986, 1989, 1996, 1998, 2000, 2006, 2009, 2011, and
603 2012) are marked. (b) Niño3.4 index (shading, NOAA-ERv4: $^{\circ}\text{C}$) and Texas precipitation
604 (green: GPCP; cyan: ERA-I: $\text{mm}\cdot\text{month}^{-1}$) from January 1979 to December 2015, each of the
605 10 EHWs is indicated by black triangle. (c) Mx2t (shading) and its 95th percent threshold
606 (curve) at a site within Texas State, based on ERA-I product.

607 **Figure 3.** (a) Probability density function of Texas summer daily Mx2t (ERA-I). The red and
608 grey curves respectively indicate the 10 EHW years and the rest 27 years; the blue and cyan
609 curves respectively indicate the early period (1980–1997) and the late period (1998–2015). (b)
610 Diagrams between Texas summer Mx2t and EHWD. The red dots indicate the 10 EHW years.

611 **Figure 4.** Composites of SSTA (NOAA-ERv4: $^{\circ}\text{C}$) and H200 (ERA-I: m) for the antecedent
612 winter (DJF), antecedent spring (MAM) and simultaneous summer (JJA) with respect to the
613 10 EHW years. Shading denotes the areas that exceed the 90% confidence level. The Texas
614 area is marked in the right panels by a red box.

615 **Figure 5.** Composites of precipitation (shading, GPCP, unit: $\text{mm}\cdot\text{month}^{-1}$) and vertical

616 integrals of water vapor flux (vector, ERA-I, units: $\text{kg}\cdot\text{m}^{-1}\cdot\text{s}^{-1}$) for the antecedent spring (a)
617 and the simultaneous summer (b) with respect to the 10 EHW years. For precipitation, the
618 negative anomalies exceeding the 95% confidence level is stippled.

619 **Figure 6.** (a) the Pacific zonal SST gradient (PZSSTG) and (b) -Niño3.4 index, which have
620 been normalized. (c) and (e) show the diagrams between normalized Texas spring and
621 summer precipitations and the PZSSTG, respectively, while (d) and (f) display the diagrams
622 between normalized Texas spring and summer precipitations and the -Niño3.4, respectively.
623 The red dots/black crosses in (e) and (f) indicate the 2011 EHW case.

624 **Figure 7.** Same as Fig. 6, except for (a)–(b) normalized Texas soil water (blue: CPC; cyan:
625 ERA-I) and (c)–(d) normalized Texas Mx2t (purple: NCEP2; green: ERA-I). Soil water is
626 averaged over 0–200 cm layers. Correlation coefficients between Texas soil water from CPC
627 and PZSSTG/-Niño3.4 from NOAA-ERAv4 are given in (c)–(f).

628 **Figure 8.** Heterogeneous correlation coefficients (CC) for the first SVD mode between
629 outgoing longwave radiation (OLR) and EHWD. (a) CCs of OLR (shading) and SST (contour)
630 with the time series of EHWD. (b) CC of EHWD with the time series of OLR. The explained
631 covariance is printed at the headlines. (c) Normalized OLR (blue) and EHWD (red) time
632 series and the linear trend in OLR (black). The square box in (a) outline the western Pacific
633 (WP) domain (100°E – 135°W , 15°S – 20°N) used to define a WP OLR index next.

634 **Figure 9.** (a) Normalized WP OLR index (black curve, downward trend indicated by blue
635 line) and NOAA-ERv4 PZSSTG (bar, upward trend indicated by red line). (b) Regression of
636 spring H200 (ERA-I, units: m) onto -WP OLR index. (c) Same as (b), except for removing

637 the Niño3.4 signals from the WP OLR index. Shadings in (b) and (c) indicate the areas
638 exceeding the 95% confidence level. Letters L and H denote the anomalous low- and
639 high-pressure centers, respectively.

640 **Figure 10.** Normalized precipitation (a) and 2-m air temperature (T2m) (b) over Texas area
641 during the boreal summer, which are obtained from each AGCM simulation (color solid) and
642 observations (blue dashed). The observational precipitation (GPCC) and Mx2t (ERA-I) have
643 been multiplied by a factor of 0.5 to match the model simulations. (c) Diagrams between
644 normalized Texas precipitation and T2m from each AGCM simulation (dots). The
645 multi-model mean (MMM) is indicated by black curves in (a)–(b) and black crosses in (c).

646 **Figure 11.** Left panels: Regression maps of the AGCM H200 (MAM, unit: m) onto PZSSTG
647 (DJF). Right panels: Similar to the left, except for removing the Niño3.4 signals from the
648 PZSSTG by linear regression. Areas exceeding the 95% confidence level are shaded.

649 **Figure 12.** Left panels: Diagrams between DJF PZSSTG and the following season's
650 precipitation and T2m over the Texas area, where these values have been normalized. Right
651 panels: Similar to the left, except for the -Niño3.4 index. Each model is marked by one
652 specific color, and the MMM is indicated by black crosses. Correlation coefficients of MMM
653 precipitation and T2m with the PZSSTG/-Niño3.4 index are given.

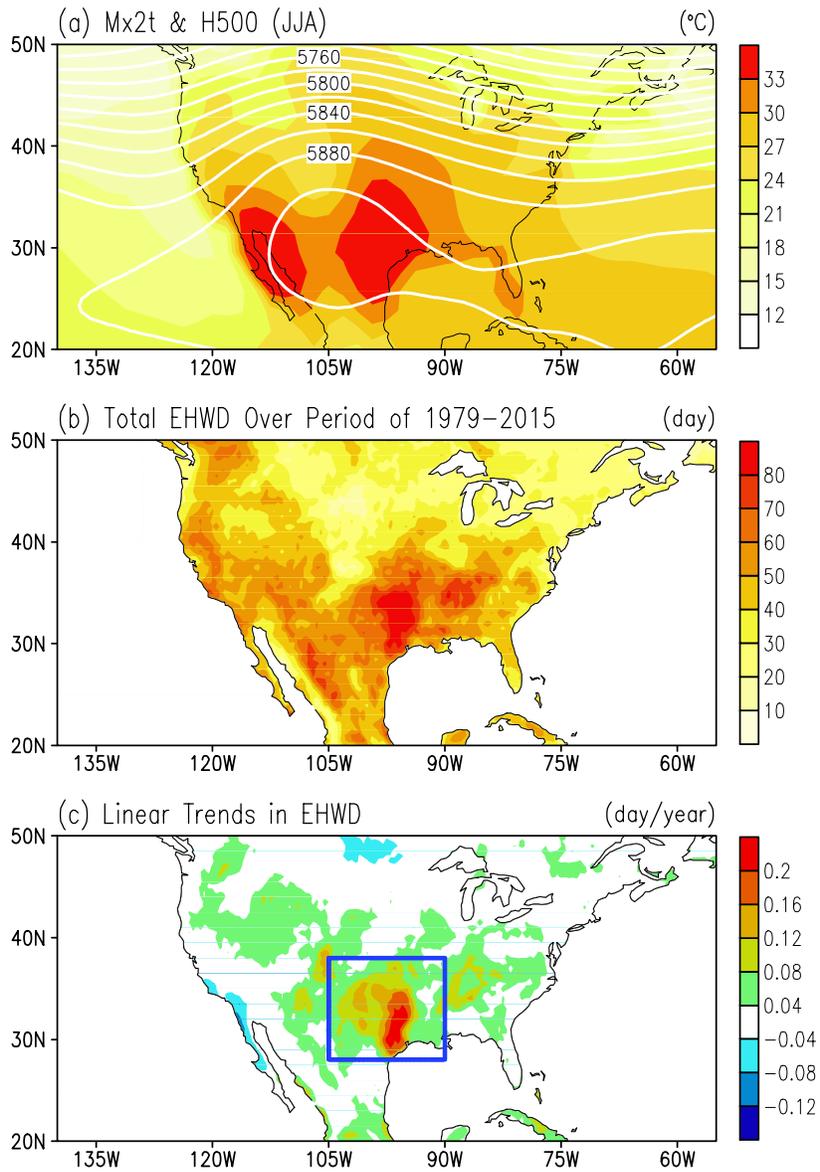
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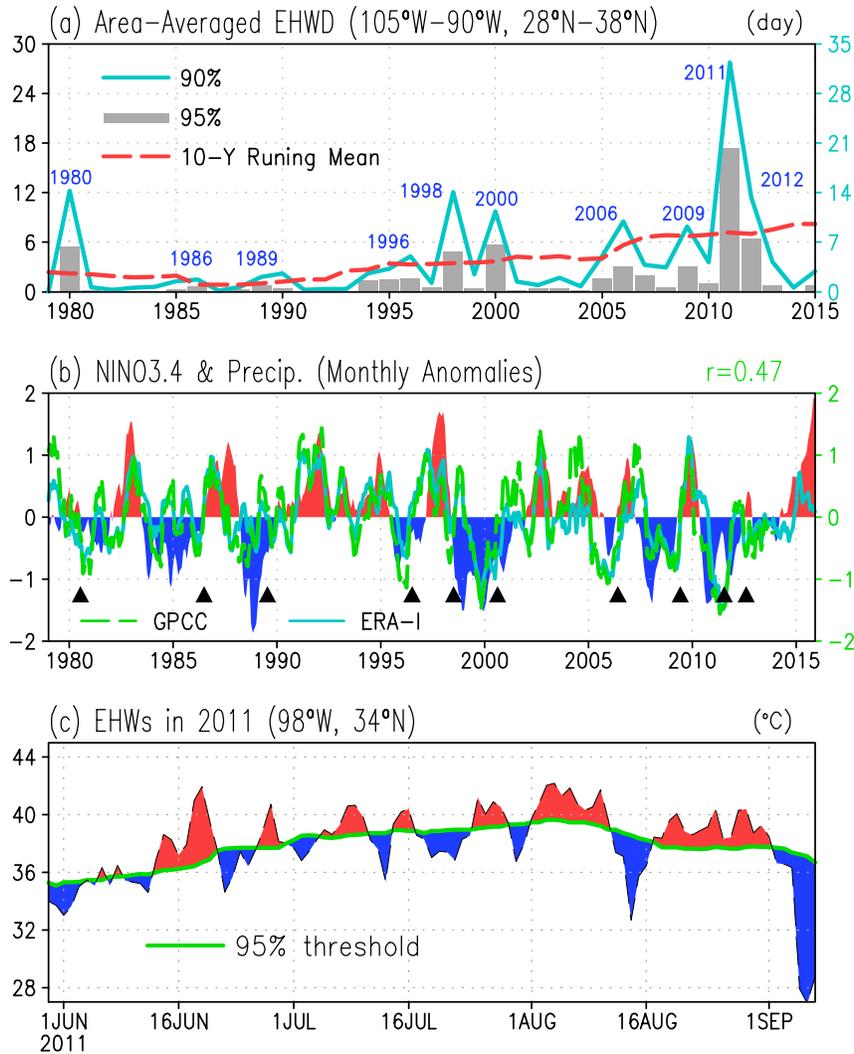
658 **List of Figures**



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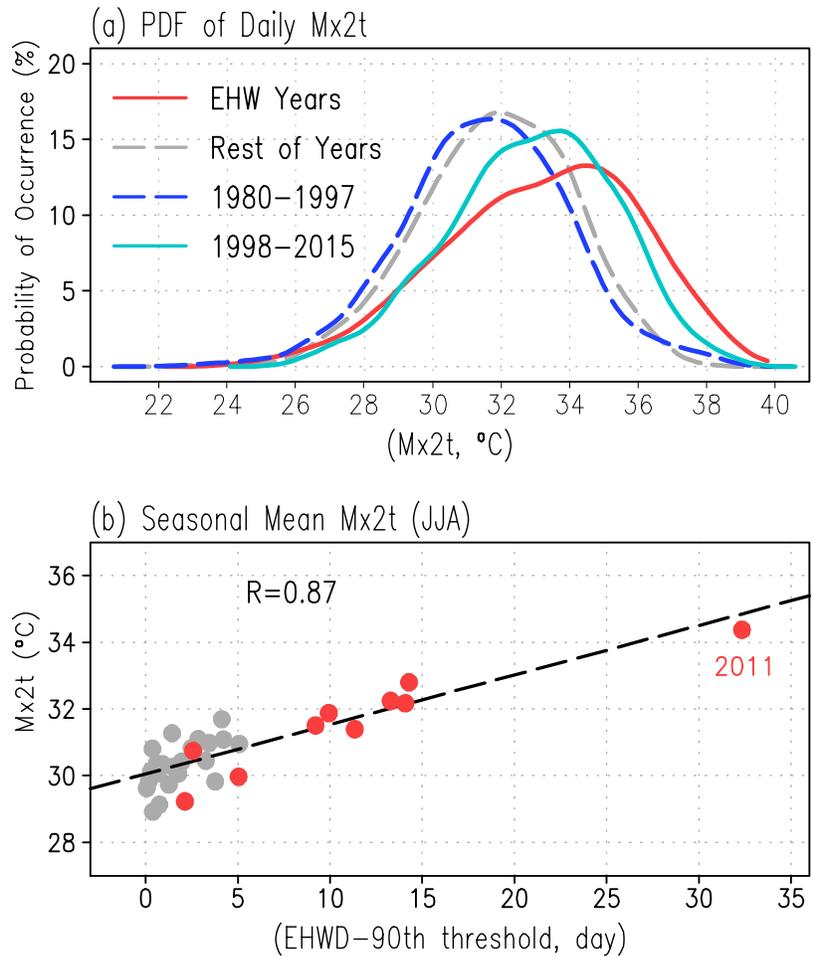
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 669 2012) are marked. (b) Niño3.4 index (shading, NOAA-ERv4: °C) and Texas precipitation
 670 (green: GPCC; cyan: ERA-I, unit: mm·month⁻¹) from January 1979 to December 2015, each
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 672 threshold (curve) at a site within Texas State, based on ERA-I product.



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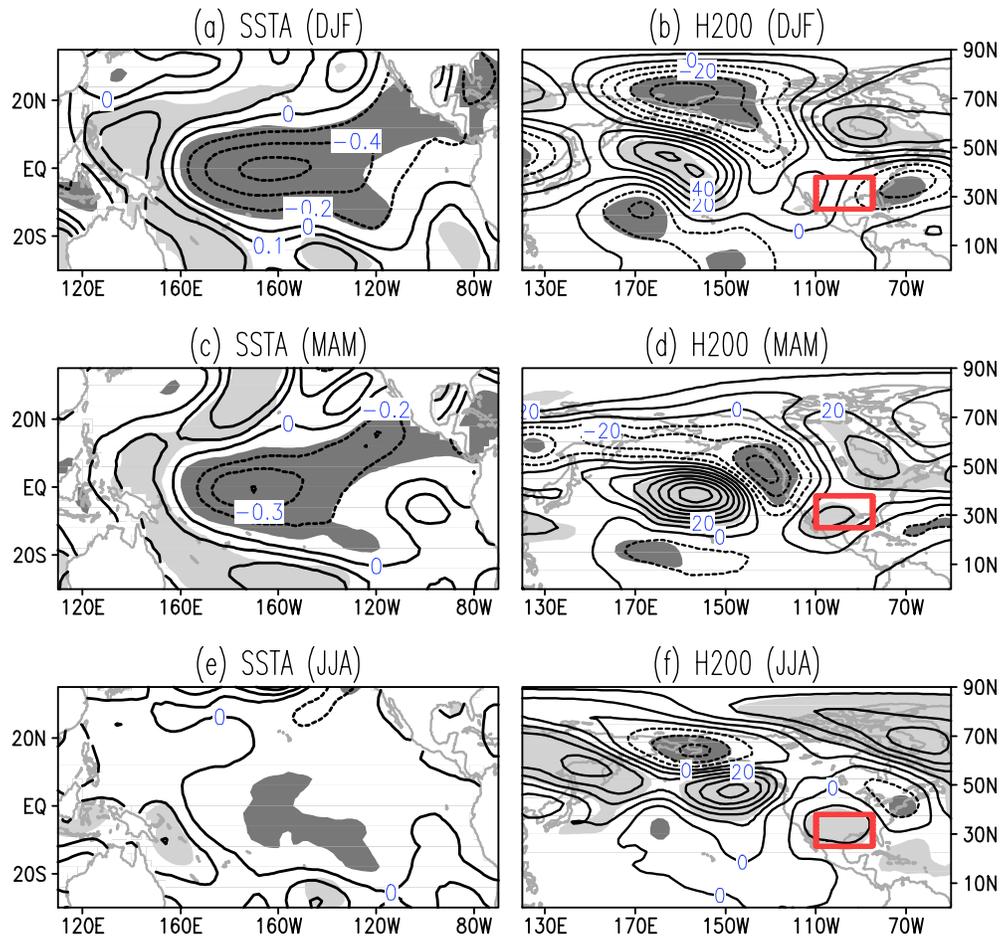
676 curves respectively indicate the early period (1980–1997) and the late period (1998–2015). (b)

677 Diagrams between Texas summer Mx2t and EHWD. The red dots indicate the 10 EHW years.

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682 **Figure 4.** Composites of SSTA (NOAA-ERv4: °C) and H200 (ERA-I: m) for the antecedent
 683 winter (DJF), antecedent spring (MAM) and simultaneous summer (JJA) with respect to the
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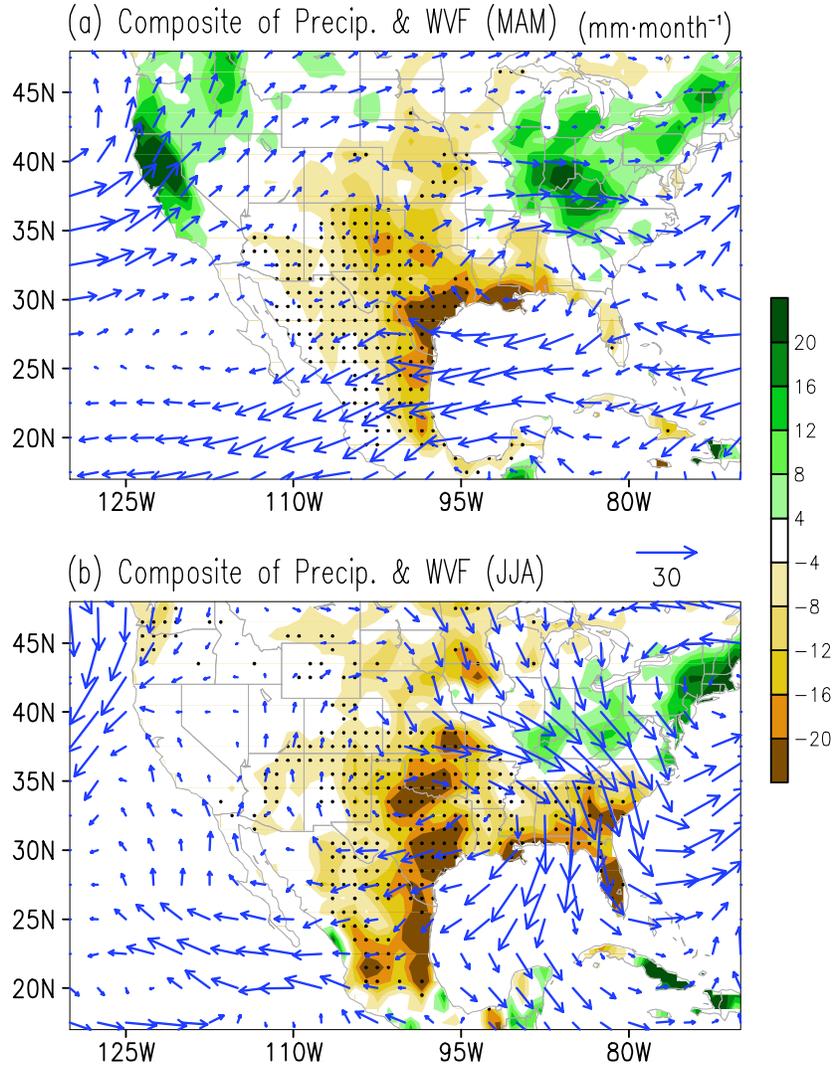
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693 integrals of water vapor flux (vector, ERA-I, unit: kg·m⁻¹·s⁻¹) for the antecedent spring (a) and

694 the simultaneous summer (b) with respect to the 10 EHW years. For precipitation, the

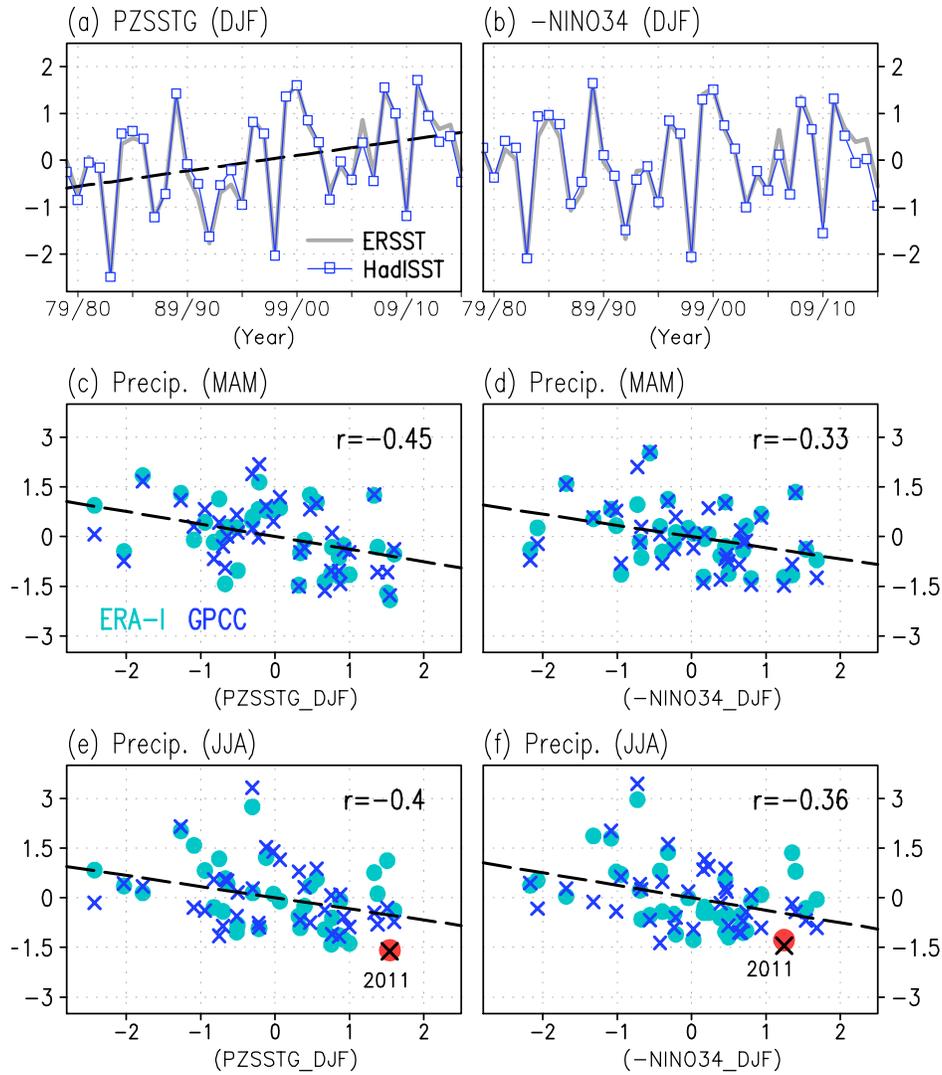
695 negative anomalies exceeding the 95% confidence level is stippled.

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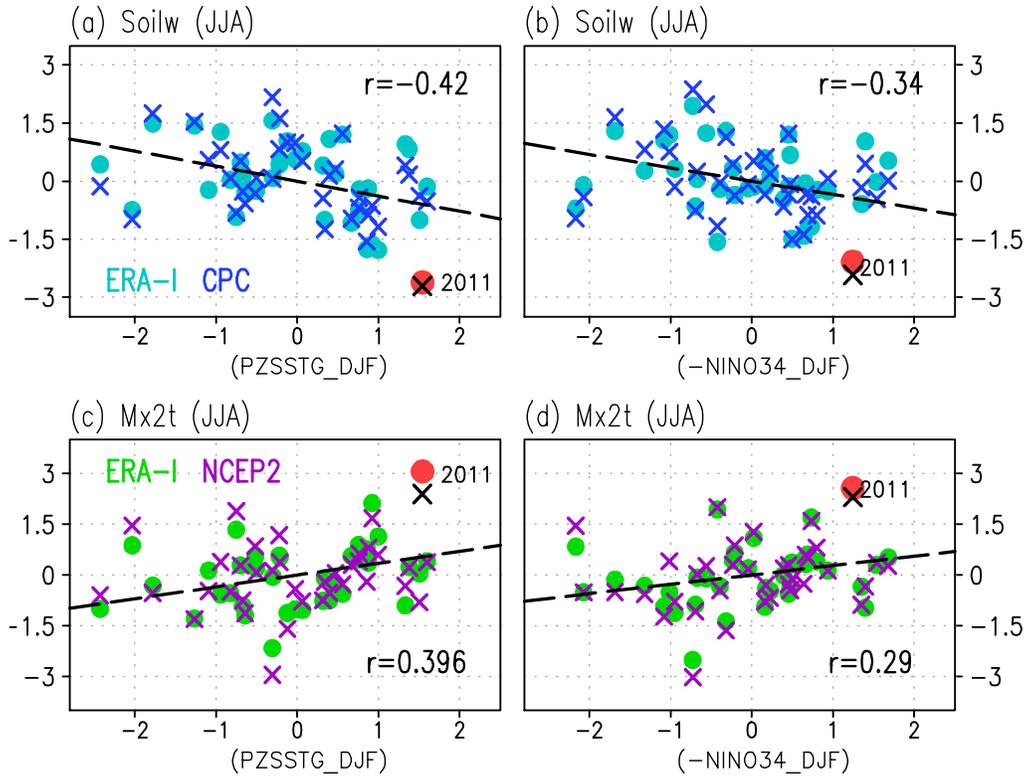
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701 **Figure 6.** (a) the Pacific zonal SST gradient (PZSSTG) and (b) -Niño3.4 index, which have
 702 been normalized. (c) and (e) show the diagrams between normalized Texas spring and
 703 summer precipitations and the PZSSTG, respectively, while (d) and (f) display the diagrams
 704 between normalized Texas spring and summer precipitations and the -Niño3.4, respectively.
 705 The red dots/black crosses in (e) and (f) indicate the 2011 EHW case.

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707

708 **Figure 7.** Same as Fig. 6, except for (a)–(b) normalized Texas soil water (blue: CPC; cyan:

709 ERA-I) and (c)–(d) normalized Texas Mx2t (purple: NCEP2; green: ERA-I). Soil water is

710 averaged over 0–200 cm layers. Correlation coefficients between Texas soil water from CPC

711 and PZSSTG/-Niño3.4 from NOAA-ERA v4 are given in (c)–(f).

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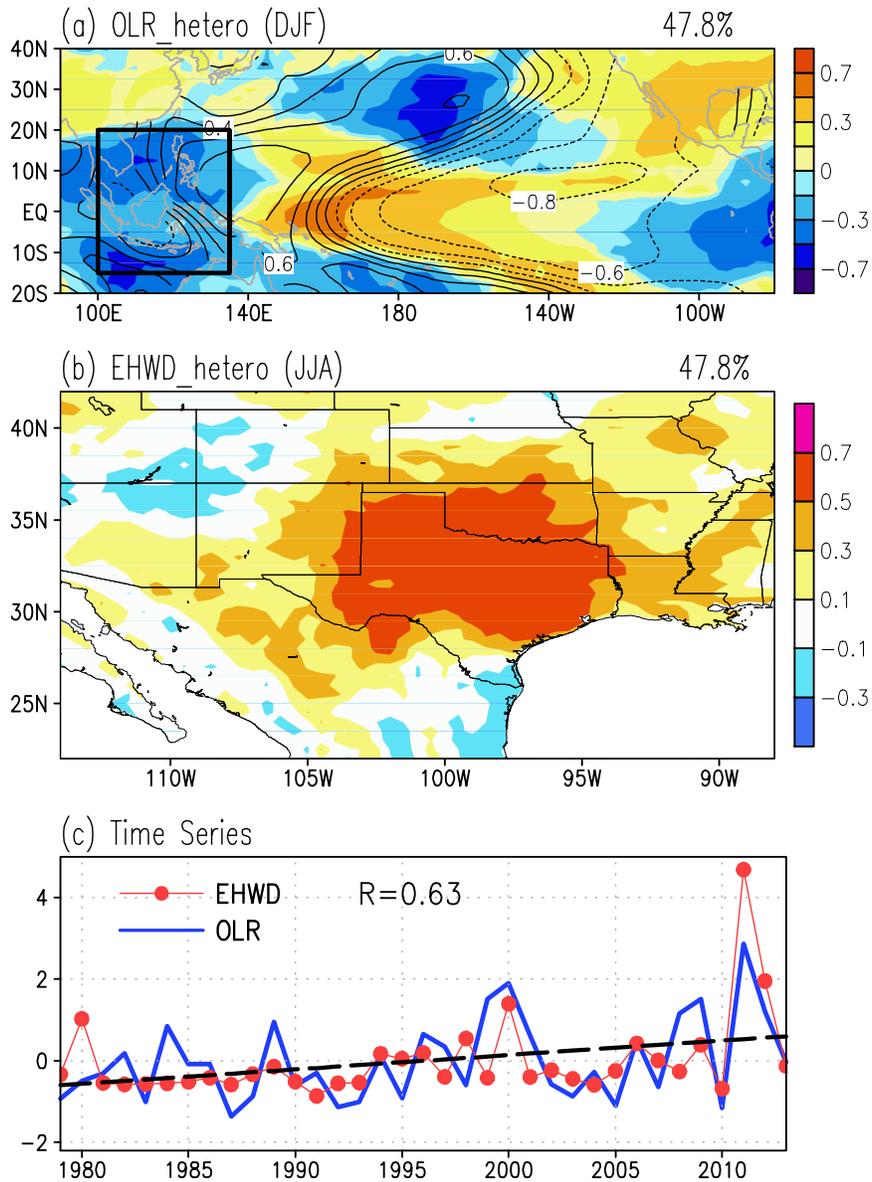
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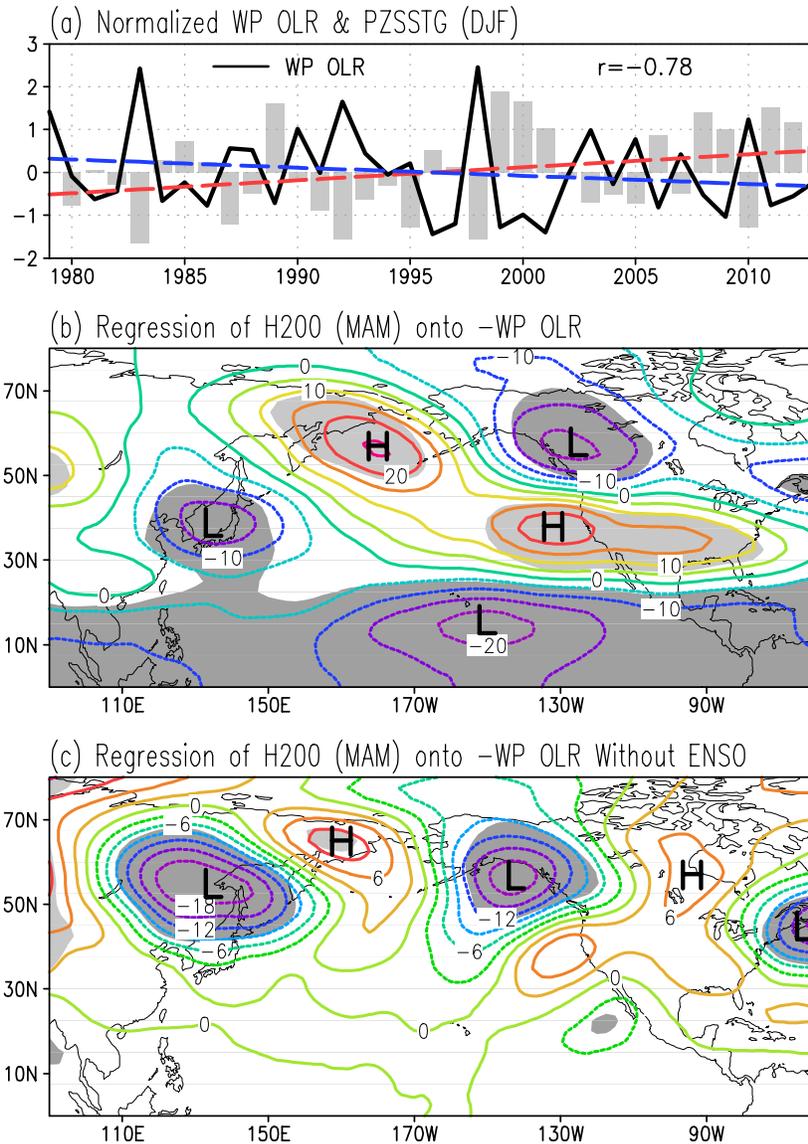
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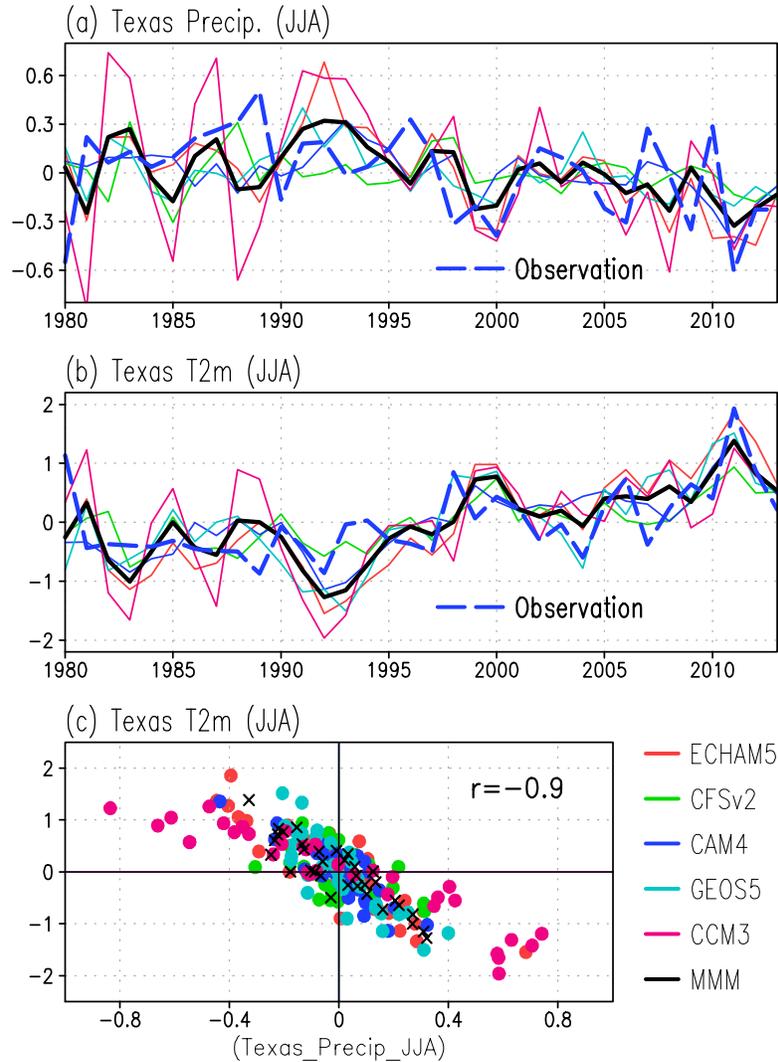
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720 **Figure 8.** Heterogeneous correlation coefficients (CC) for the first SVD mode between
 721 outgoing longwave radiation (OLR) and EHWD. (a) CCs of OLR (shading) and SST (contour)
 722 with the time series of EHWD. (b) CC of EHWD with the time series of OLR. The explained
 723 covariance is printed at the headlines. (c) Normalized OLR (blue) and EHWD (red) time
 724 series and the linear trend in OLR (black). The square box in (a) outline the western Pacific
 725 (WP) domain (100°E–135°W, 15°S–20°N) used to define a WP OLR index next.



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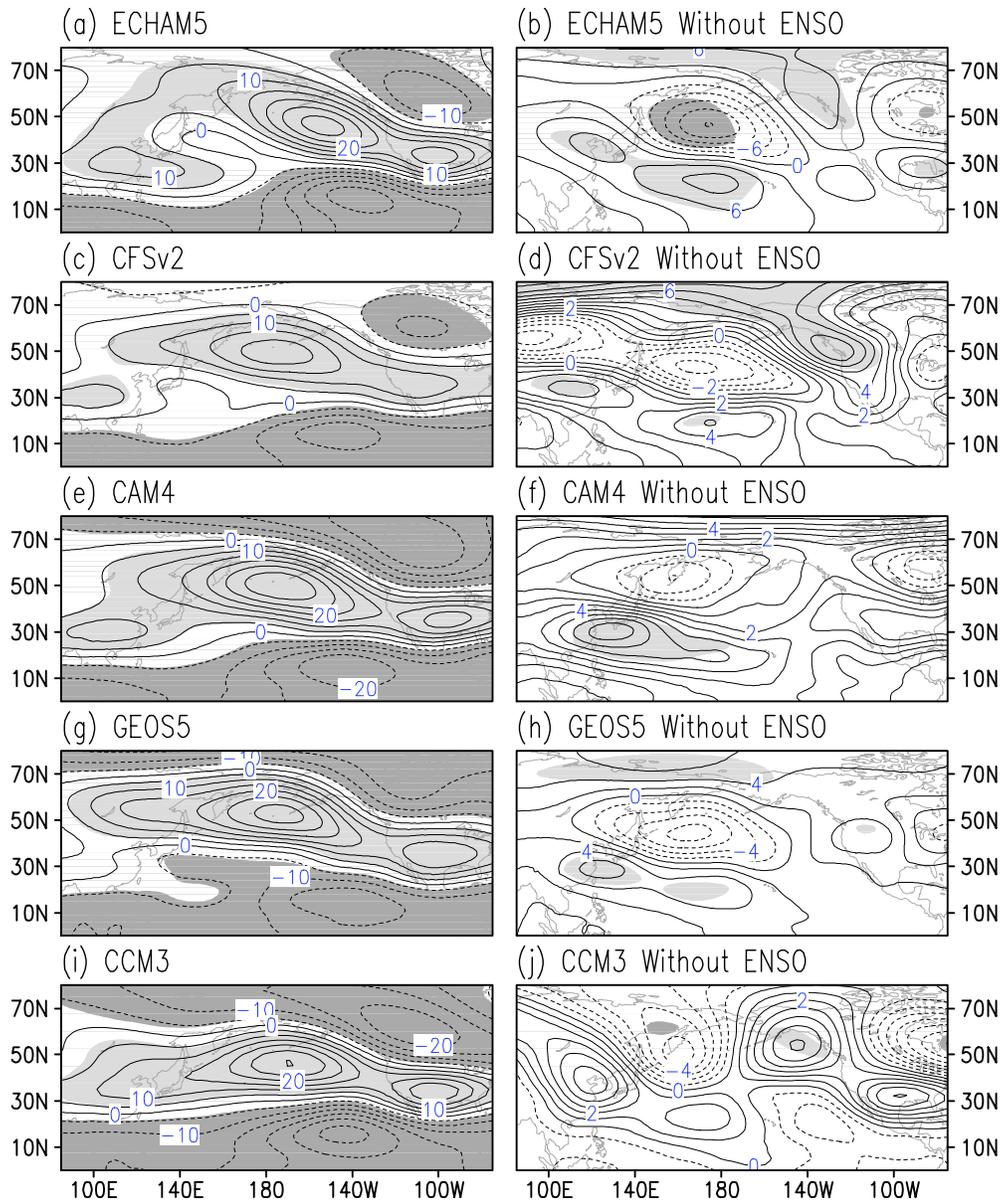
727 **Figure 9.** (a) Normalized WP OLR index (black curve, downward trend indicated by blue
 728 line) and NOAA-ERv4 PZSSTG (bar, upward trend indicated by red line). (b) Regression of
 729 spring H200 (ERA-I, units: m) onto -WP OLR index. (c) Same as (b), except for removing
 730 the Niño3.4 signals from the WP OLR index. Shadings in (b) and (c) indicate the areas
 731 exceeding the 95% confidence level. Letters L and H denote the anomalous low- and
 732 high-pressure centers, respectively.



733

734 **Figure 10.** Normalized precipitation (a) and 2-m air temperature (T2m) (b) over Texas area
 735 during the boreal summer, which are obtained from each AGCM simulation (color solid) and
 736 observations (blue dashed). The observational precipitation (GPCC) and Mx2t (ERA-I) have
 737 been multiplied by a factor of 0.5 to match the model simulations. (c) Diagrams between
 738 Texas precipitation and T2m from each AGCM simulation (dots). The multi-model ensemble
 739 mean (MMM) is indicated by black curves in (a)–(b) and black crosses in (c).

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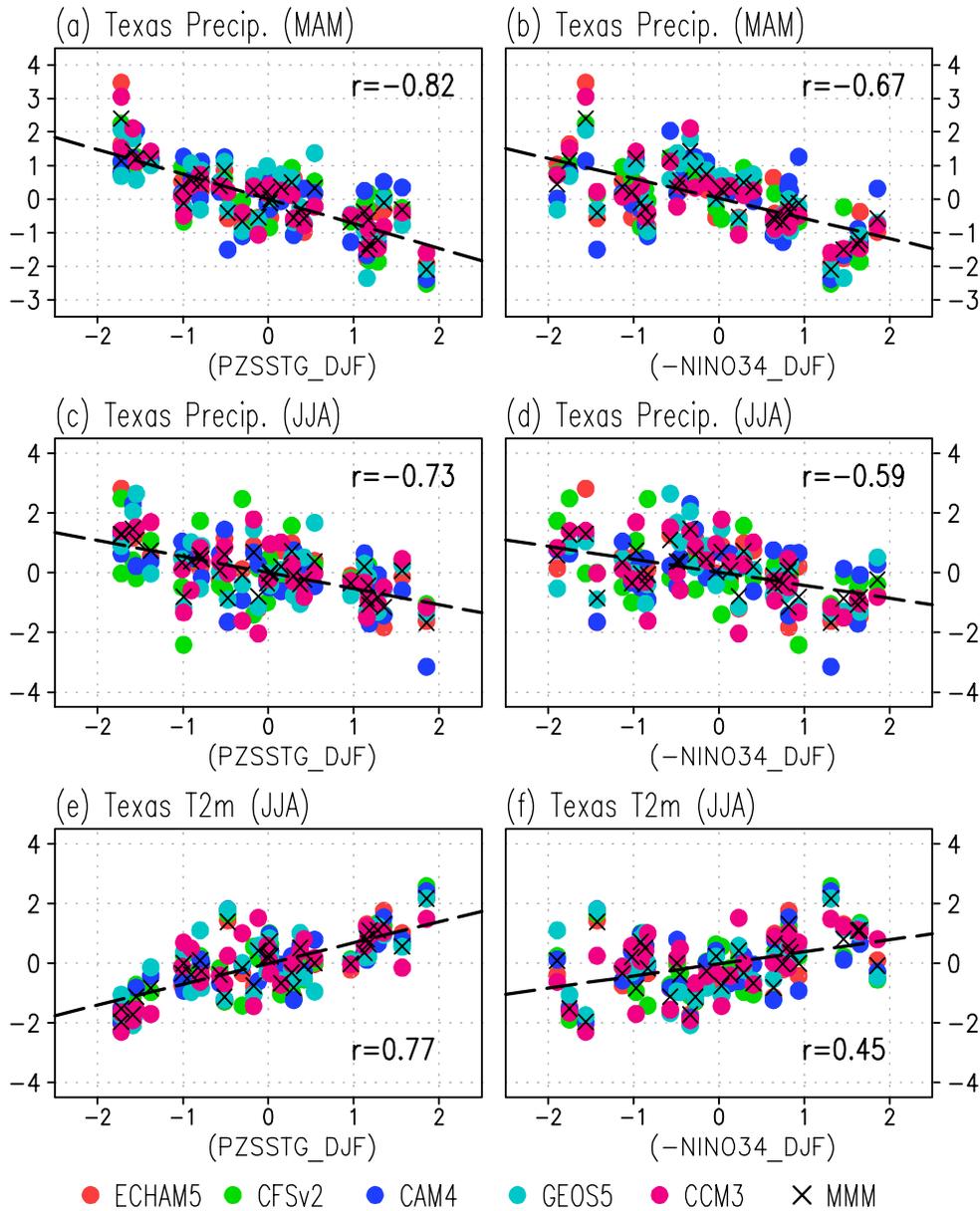
742 **Figure 11.** Left panels: Regression maps of the AGCM H200 (MAM, unit: m) onto PZSSTG

743 (DJF). Right panels: Similar to the left, except for removing the Niño3.4 signals from the

744 PZSSTG by linear regression. Areas exceeding the 95% confidence level are shaded.

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746



748 **Figure 12.** Left panels: Diagrams between DJF PZSSTG and the following season's
 749 precipitation and T2m over the Texas area, where these values have been normalized. Right
 750 panels: Similar to the left, except for the -Niño3.4 index. Each model is marked by one
 751 specific color, and the MMM is indicated by black crosses. Correlation coefficients of MMM
 752 precipitation and T2m with the PZSSTG/-Niño3.4 index are given.

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