Lecture 9

Inexact Theories

Syllabus

Lecture 01 Describing Inverse Problems Probability and Measurement Error, Part 1 Lecture 02 Probability and Measurement Error, Part 2 Lecture 03 Lecture 04 The L₂ Norm and Simple Least Squares A Priori Information and Weighted Least Squared Lecture 05 **Resolution and Generalized Inverses** Lecture 06 Lecture 07 Backus-Gilbert Inverse and the Trade Off of Resolution and Variance Lecture 08 The Principle of Maximum Likelihood Lecture 09 **Inexact Theories** Lecture 10 Nonuniqueness and Localized Averages Vector Spaces and Singular Value Decomposition Lecture 11 Lecture 12 Equality and Inequality Constraints Lecture 13 L_1 , L_{∞} Norm Problems and Linear Programming Lecture 14 Nonlinear Problems: Grid and Monte Carlo Searches Nonlinear Problems: Newton's Method Lecture 15 Lecture 16 Nonlinear Problems: Simulated Annealing and Bootstrap Confidence Intervals Lecture 17 **Factor Analysis** Varimax Factors, Empirical Orthogonal Functions Lecture 18 Lecture 19 Backus-Gilbert Theory for Continuous Problems; Radon's Problem Lecture 20 Linear Operators and Their Adjoints Lecture 21 Fréchet Derivatives Lecture 22 Exemplary Inverse Problems, incl. Filter Design Lecture 23 Exemplary Inverse Problems, incl. Earthquake Location Lecture 24 Exemplary Inverse Problems, incl. Vibrational Problems

Purpose of the Lecture

Discuss how an inexact theory can be represented

Solve the inexact, linear Gaussian inverse problem

Use maximization of relative entropy as a guiding principle for solving inverse problems

Introduce F-test as way to determine whether one solution is "better" than another

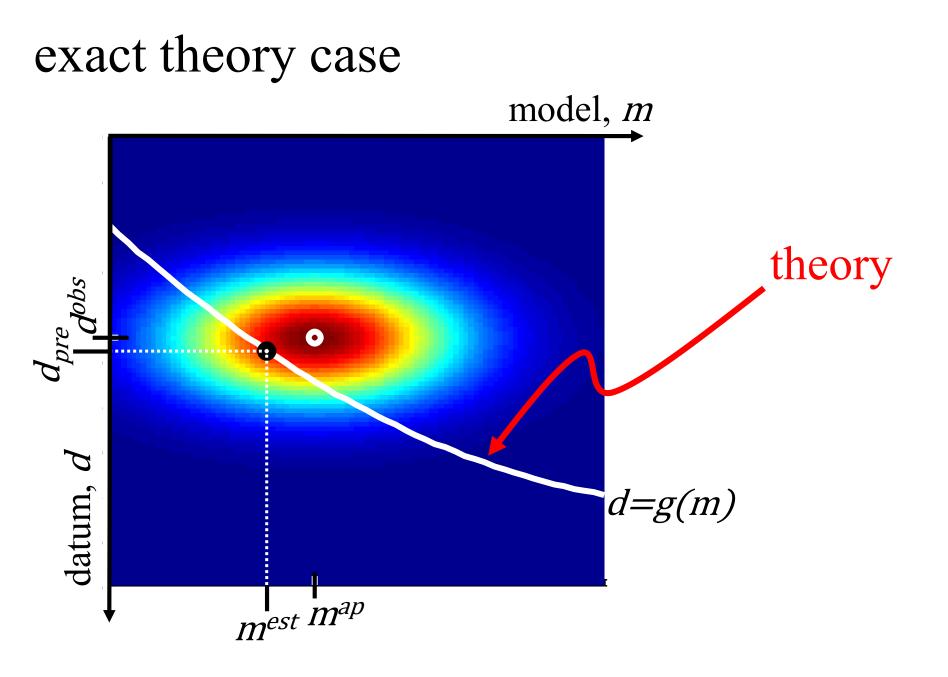
Part 1

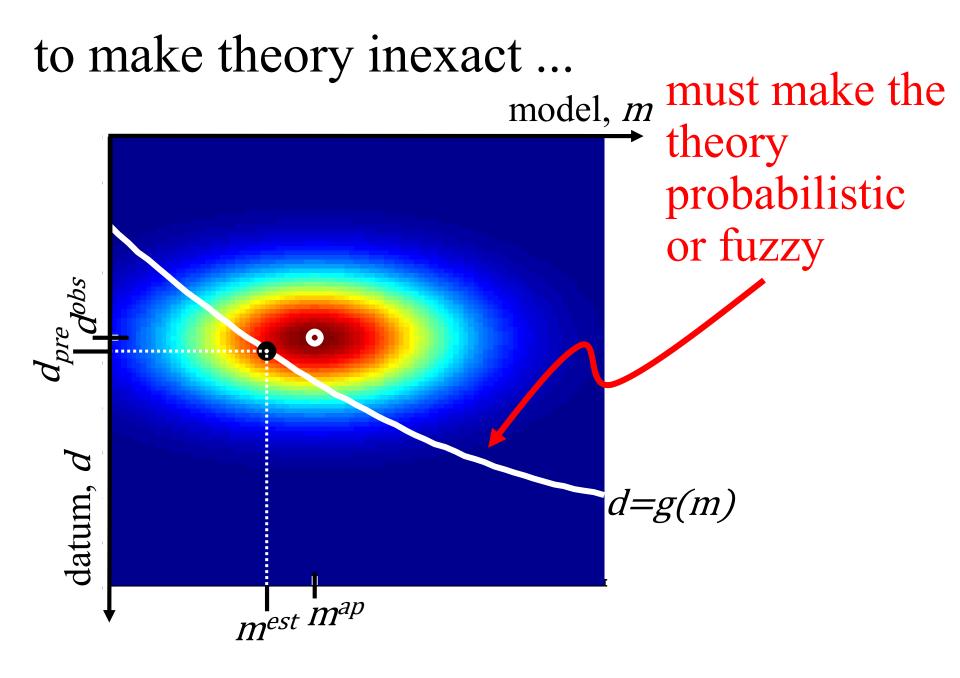
How Inexact Theories can be Represented

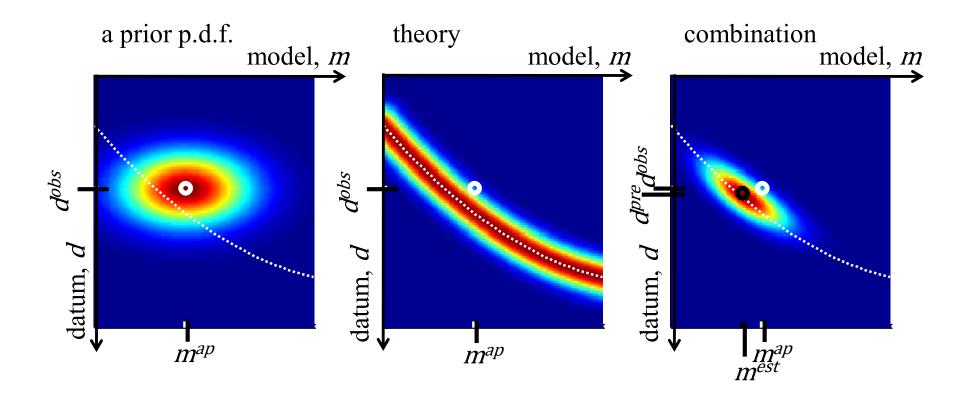
How do we generalize the case of

an exact theory

to one that is inexact?







how do you *combine* two probability density functions ?

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so that the information in them is combined ...

desirable properties

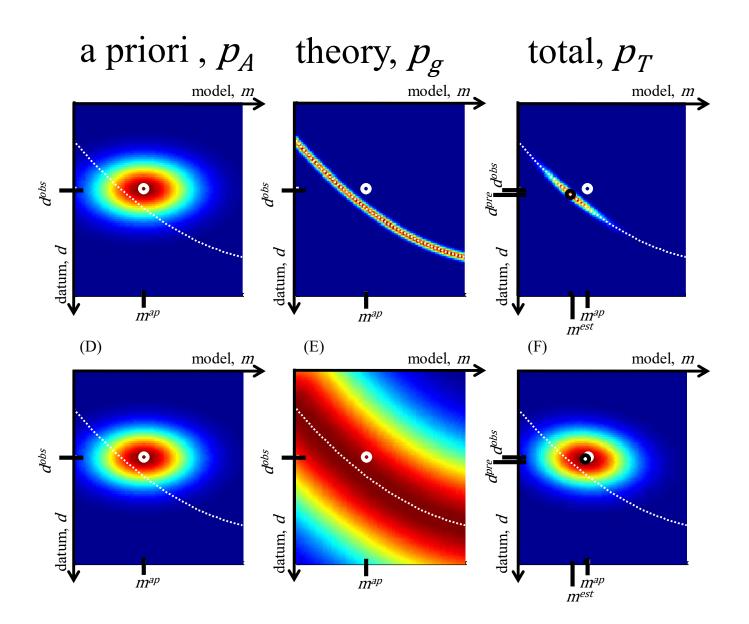
order shouldn't matter

combining something with the null distribution should leave it unchanged

combination should be invariant under change of variables

Answer

 $p_3 = \frac{p_1 p_2}{p_N}$



"solution to inverse problem" maximum likelihood point of

$p_T(\mathbf{m}, \mathbf{d}) = p_A(\mathbf{m}, \mathbf{d}) p_g(\mathbf{m}, \mathbf{d})$ (with $p_N \propto constant$)

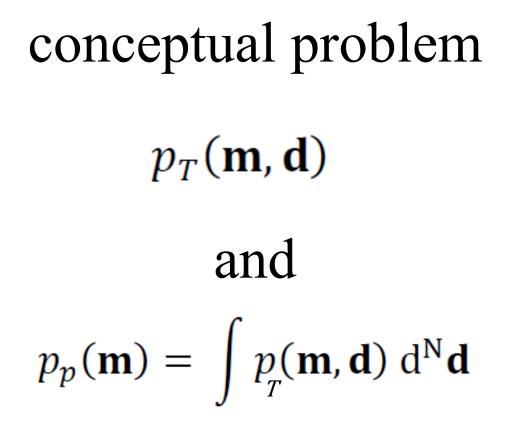
simultaneously gives m^{est} and d^{pre}

$p_T(\mathbf{m}, \mathbf{d})$

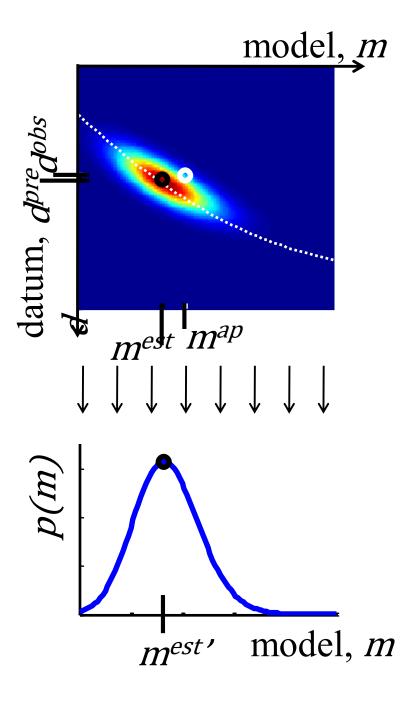
probability that the estimated model parameters are near **m** and the predicted data are near **d**

$$p_p(\mathbf{m}) = \int p_T(\mathbf{m}, \mathbf{d}) \, \mathrm{d}^{\mathrm{N}} \mathbf{d}$$

probability that the estimated model parameters are near **m** irrespective of the value of the predicted data



do not necessarily have maximum likelihood points at the same value of **m**



illustrates the problem in defining a **definitive solution** to an inverse problem

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fortunately if all distributions are Gaussian the two points are the same

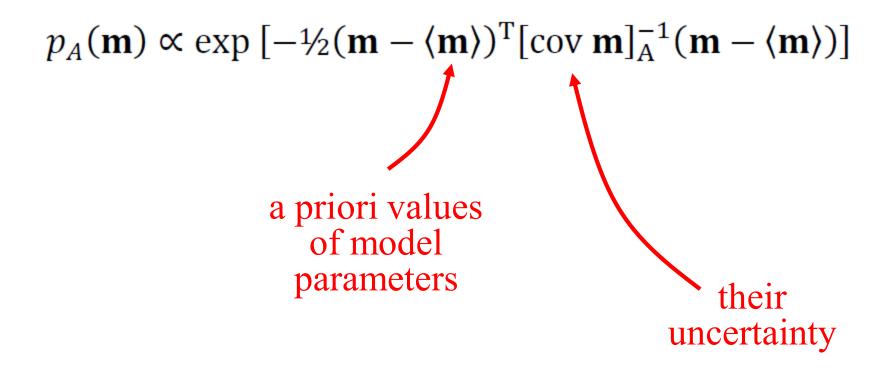
Part 2

Solution of the inexact linear Gaussian inverse problem

Gaussian a priori information

 $p_A(\mathbf{m}) \propto \exp\left[-\frac{1}{2}(\mathbf{m} - \langle \mathbf{m} \rangle)^{\mathrm{T}} [\operatorname{cov} \mathbf{m}]_{\mathrm{A}}^{-1}(\mathbf{m} - \langle \mathbf{m} \rangle)\right]$

Gaussian a priori information



Gaussian observations

$$p_A(\mathbf{d}) \propto \exp\left[-\frac{1}{2}\left(\mathbf{d} - \mathbf{d}^{\text{obs}}\right)^{\text{T}} [\operatorname{cov} \mathbf{d}]^{-1} \left(\mathbf{d} - \mathbf{d}^{\text{obs}}\right)\right]$$

Gaussian observations

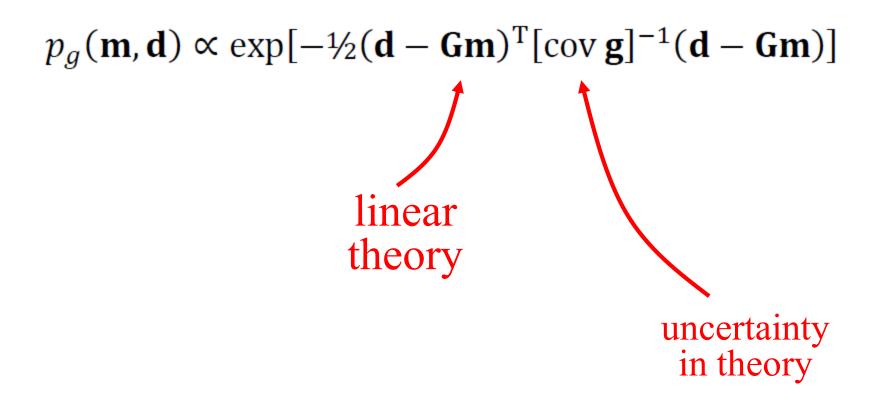
$$p_{A}(\mathbf{d}) \propto \exp\left[-\frac{1}{2}\left(\mathbf{d} - \mathbf{d}^{\text{obs}}\right)^{\text{T}}\left[\operatorname{cov} \mathbf{d}\right]^{-1}\left(\mathbf{d} - \mathbf{d}^{\text{obs}}\right)\right]$$

observed
data
measurement
error

Gaussian theory

$p_g(\mathbf{m}, \mathbf{d}) \propto \exp[-\frac{1}{2}(\mathbf{d} - \mathbf{G}\mathbf{m})^{\mathrm{T}}[\operatorname{cov} \mathbf{g}]^{-1}(\mathbf{d} - \mathbf{G}\mathbf{m})]$

Gaussian theory



mathematical statement of problem

find (m,d) that maximizes

 $p_{\mathrm{T}}(\mathbf{m},\mathbf{d}) = p_{\mathrm{A}}(\mathbf{m}) p_{\mathrm{A}}(\mathbf{d}) p_{\mathrm{g}}(\mathbf{m},\mathbf{d})$

and, along the way, work out the form of $p_{\rm T}({\bf m},{\bf d})$

notational simplification

group **m** and **d** into single vector $\mathbf{x} = [\mathbf{d}^T, \mathbf{m}^T]^T$

group $[\operatorname{cov} \mathbf{m}]_{A}$ and $[\operatorname{cov} \mathbf{d}]_{A}$ into single matrix $[\operatorname{cov} \mathbf{x}] = \begin{bmatrix} [\operatorname{cov} \mathbf{d}] & 0\\ 0 & [\operatorname{cov} \mathbf{m}]_{A} \end{bmatrix}$

write d-Gm=0 as Fx=0 with F=[I, -G]

after much algebra, we find $p_{\rm T}({\bf x})$ is a Gaussian distribution with mean $\mathbf{x}^* = \left\{ \mathbf{I} - [\operatorname{cov} \mathbf{x}] \mathbf{F}^{\mathrm{T}} [[\operatorname{cov} \mathbf{g}] + \mathbf{F} [\operatorname{cov} \mathbf{x}] \mathbf{F}^{\mathrm{T}}]^{-1} \mathbf{F} \right\} \langle \mathbf{x} \rangle$ and variance $[\operatorname{cov} \mathbf{x}^*] = \left\{ \mathbf{I} - [\operatorname{cov} \mathbf{x}] \mathbf{F}^{\mathrm{T}} [[\operatorname{cov} \mathbf{g}] + \mathbf{F} [\operatorname{cov} \mathbf{x}] \mathbf{F}^{\mathrm{T}} \right\}^{-1} \mathbf{F} \left\{ [\operatorname{cov} \mathbf{x}] \mathbf{F} \left\{ [\operatorname$

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after pulling m^{est} out of x*

$$\mathbf{m}^{\text{est}} = \langle \mathbf{m} \rangle + \mathbf{G}^{-g} (\mathbf{d}^{\text{obs}} - \mathbf{G} \langle \mathbf{m} \rangle) = \mathbf{G}^{-g} \mathbf{d}^{\text{obs}} + [\mathbf{I} - \mathbf{R}] \langle \mathbf{m} \rangle$$

with $\mathbf{G}^{-g} = [\operatorname{cov} \mathbf{m}]_{A}\mathbf{G}^{T}\{[\operatorname{cov} \mathbf{d}] + [\operatorname{cov} \mathbf{g}] + \mathbf{G}[\operatorname{cov} \mathbf{m}]_{A}\mathbf{G}^{T}\}^{-1}$

after pulling
$$\mathbf{m}^{\text{est}}$$
 out of \mathbf{x}^*
 $\mathbf{m}^{\text{est}} = \langle \mathbf{m} \rangle + \mathbf{G}^{-g} (\mathbf{d}^{\text{obs}} - \mathbf{G} \langle \mathbf{m} \rangle) = \mathbf{G}^{-g} \mathbf{d}^{\text{obs}} + [\mathbf{I} - \mathbf{R}] \langle \mathbf{m} \rangle$
with $\mathbf{G}^{-g} = [\operatorname{cov} \mathbf{m}]_A \mathbf{G}^T \{[\operatorname{cov} \mathbf{d}] + [\operatorname{cov} g] + \mathbf{G} [\operatorname{cov} \mathbf{m}]_A \mathbf{G}^T \}^{-1}$

reminiscent of **G**^T(**GG**^T)⁻¹ minimum length solution

after pulling
$$\mathbf{m}^{est}$$
 out of \mathbf{x}^*
 $\mathbf{m}^{est} = \langle \mathbf{m} \rangle + \mathbf{G}^{-g} (\mathbf{d}^{obs} - \mathbf{G} \langle \mathbf{m} \rangle) = \mathbf{G}^{-g} \mathbf{d}^{obs} + [\mathbf{I} - \mathbf{R}] \langle \mathbf{m} \rangle$
with $\mathbf{G}^{-g} = [\operatorname{cov} \mathbf{m}]_A \mathbf{G}^T [[\operatorname{cov} \mathbf{d}] + [\operatorname{cov} \mathbf{g}] + \mathbf{G} [[\operatorname{cov} \mathbf{m}]_A \mathbf{G}^T]^{-1}$

error in theory adds to error in data

after pulling m^{est} out of x*

$$\mathbf{m}^{\text{est}} = \langle \mathbf{m} \rangle + \mathbf{G}^{-g} (\mathbf{d}^{\text{obs}} - \mathbf{G} \langle \mathbf{m} \rangle) = \mathbf{G}^{-g} \mathbf{d}^{\text{obs}} + [\mathbf{I} - \mathbf{R}] \langle \mathbf{m} \rangle$$

with $\mathbf{G}^{-g} = [\operatorname{cov} \mathbf{m}]_{A}\mathbf{G}^{T}\{[\operatorname{cov} \mathbf{d}] + [\operatorname{cov} \mathbf{g}] + \mathbf{G}[\operatorname{cov} \mathbf{m}]_{A}\mathbf{G}^{T}\}^{-1}$

solution depends on the values of the prior information only to the extent that the model resolution matrix is different from an identity matrix

and after algebraic manipulation

$$\mathbf{m}^{\text{est}} = \langle \mathbf{m} \rangle + \mathbf{G}^{-g} (\mathbf{d}^{\text{obs}} - \mathbf{G} \langle \mathbf{m} \rangle) = \mathbf{G}^{-g} \mathbf{d}^{\text{obs}} + [\mathbf{I} - \mathbf{R}] \langle \mathbf{m} \rangle$$

with $\mathbf{G}^{-g} = [\operatorname{cov} \mathbf{m}]_{A}\mathbf{G}^{T}\{[\operatorname{cov} \mathbf{d}] + [\operatorname{cov} \mathbf{g}] + \mathbf{G}[\operatorname{cov} \mathbf{m}]_{A}\mathbf{G}^{T}\}^{-1}$

which also equals

$$\mathbf{G}^{-\mathbf{g}} = \left(\mathbf{G}^{\mathrm{T}}([\operatorname{cov} \mathbf{d}] + [\operatorname{cov} \mathbf{g}])^{-1}\mathbf{G} + [\operatorname{cov} \mathbf{m}]_{A}^{-1}\right)^{-1}\mathbf{G}^{\mathrm{T}}([\operatorname{cov} \mathbf{d}] + [\operatorname{cov} \mathbf{g}])^{-1}$$
reminiscent of $(\mathbf{G}^{\mathrm{T}}\mathbf{G})^{-1}\mathbf{G}^{\mathrm{T}}$
least squares solution

interesting aside

weighted least squares solution

is equal to the

weighted minimum length solution

what did we learn?

for linear Gaussian inverse problem

inexactness of theory just adds to inexactness of data

Part 3

Use maximization of relative entropy as a guiding principle for solving inverse problems

from last lecture

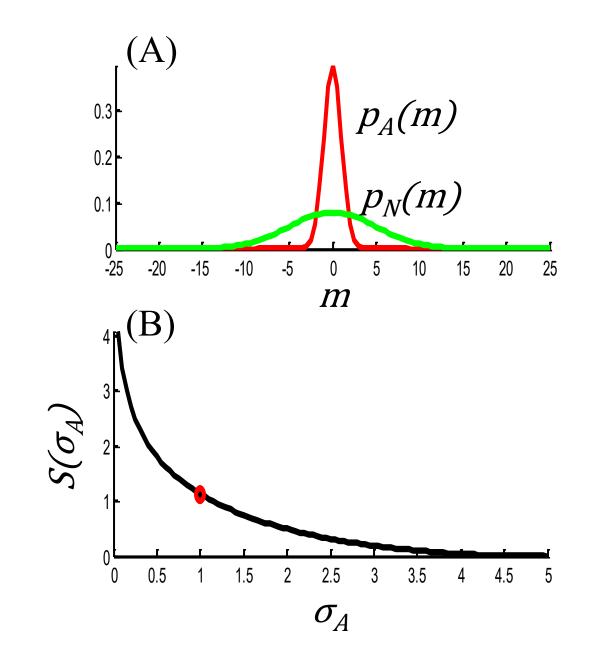
assessing the information content in $p_A(\mathbf{m})$

Do we know a little about **m** or a lot about **m** ?

Information Gain, S

$$S[p_{A}(\mathbf{m})] = \int p_{A}(\mathbf{m}) \log \left[\frac{p_{A}(\mathbf{m})}{p_{N}(\mathbf{m})}\right] d^{M}\mathbf{m}$$

-S called Relative Entropy



Principle of Maximum Relative Entropy

or if you prefer

Principle of Minimum Information Gain

find solution p.d.f. $p_{\rm T}(\mathbf{m})$ that has the largest relative entropy as compared to a priori p.d.f. $p_{\rm A}(\mathbf{m})$

or if you prefer

find solution p.d.f. $p_{\rm T}(\mathbf{m})$ that has smallest possible new information as compared to a priori p.d.f. $p_{\rm A}(\mathbf{m})$

minimize
$$S = \int p_T(\mathbf{m}) \log\left(\frac{p_T(\mathbf{m})}{p_A(\mathbf{m})}\right) d^{\mathbf{M}}m$$
 with constraints
 $\int p_T(\mathbf{m}) d^{\mathbf{M}}m = 1$ and $\int p_T(\mathbf{m}) (\mathbf{d} - \mathbf{Gm}) d^{\mathbf{M}}m = 0$

minimize
$$S = \int p_T(\mathbf{m}) \log \left(\frac{p_T(\mathbf{m})}{p_A(\mathbf{m})} \right) d^M m$$
 with constraints

$$\int p_T(\mathbf{m}) d^M m = 1 \text{ and } \int p_T(\mathbf{m}) (\mathbf{d} - \mathbf{Gm}) d^M m = 0$$

$$\sum_{\substack{\text{properly} \\ \text{normalized} \\ p.d.f.}} data \text{ is satisfied in the mean} or expected value of error is zero}$$

After minimization using Lagrange Multipliers process

 $p_{\rm T}({\bf m})$ is Gaussian with maximum likelihood point ${\bf m}^{\rm est}$ satisfying

 $\mathbf{m}^{\text{est}} - \langle \mathbf{m} \rangle = [\operatorname{cov} \mathbf{m}]_A \mathbf{G}^{\mathrm{T}} \{ \mathbf{G} [\operatorname{cov} \mathbf{m}]_A \mathbf{G}^{\mathrm{T}} \}^{-1} \{ \mathbf{d} - \mathbf{G} \langle \mathbf{m} \rangle \}$

After minimization using Lagrane Multipliers process

 $p_{\rm T}({\bf m})$ is Gaussian with maximum likelihood point ${\bf m}^{\rm est}$ satisfying

 $\mathbf{m}^{\text{est}} - \langle \mathbf{m} \rangle = [\operatorname{cov} \mathbf{m}]_A \mathbf{G}^{\mathrm{T}} \{ \mathbf{G} [\operatorname{cov} \mathbf{m}]_A \mathbf{G}^{\mathrm{T}} \}^{-1} \{ \mathbf{d} - \mathbf{G} \langle \mathbf{m} \rangle \}$

just the weighted minimum length solution

What did we learn?

Only that the Principle of Maximum Entropy is yet another way of deriving the inverse problem solutions we are already familiar with

Part 4

F-test

as way to determine whether one solution is "better" than another

Common Scenario

two different theories

solution m^{est}_{A} M_{A} model parameters prediction error E_{A}

solution m^{est}_{B} M_{B} model parameters prediction error E_{B}

Suppose $E_B < E_A$

Is B really better than A?

What if B has many more model parameters than A

 $M_R >> M_A$

Is B fitting better any surprise?

Need to against Null Hypothesis

The difference in error is due to random variation

suppose error **e** has a Gaussian p.d.f. uncorrelated uniform variance σ_d

estimate variance

$$(\sigma_d^{est})^2 = \frac{1}{\nu} \sum_{i=1}^{N} e_i^2 = \frac{E}{\nu}$$
 with $\nu = N - M$

want to known the probability density function of

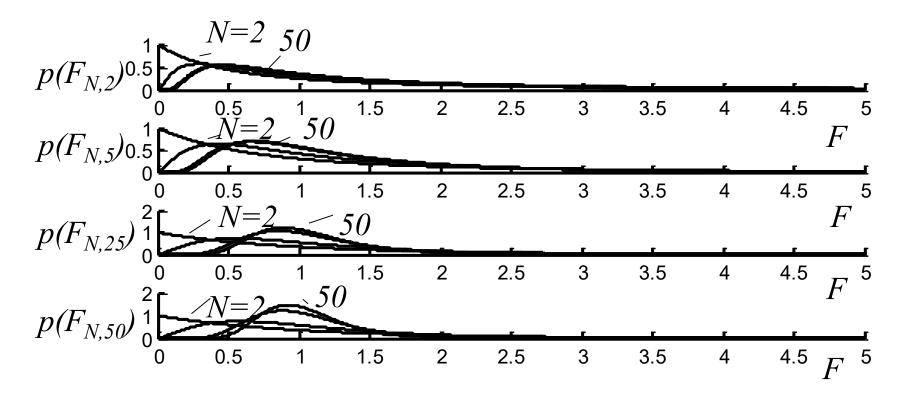
$(\sigma_{dA}^{est})^2/(\sigma_{dB}^{est})^2$

actually, we'll use the quantity

$$F(\nu_A, \nu_B) = \frac{(\sigma_{dA}^{est})^2 / (\sigma_{dA}^{true})^2}{(\sigma_{dB}^{est})^2 / (\sigma_{dA}^{true})^2}$$

which is the same, as long as the two theories that we're testing is applied to the same data

p.d.f. of F is known



as is its mean and variance

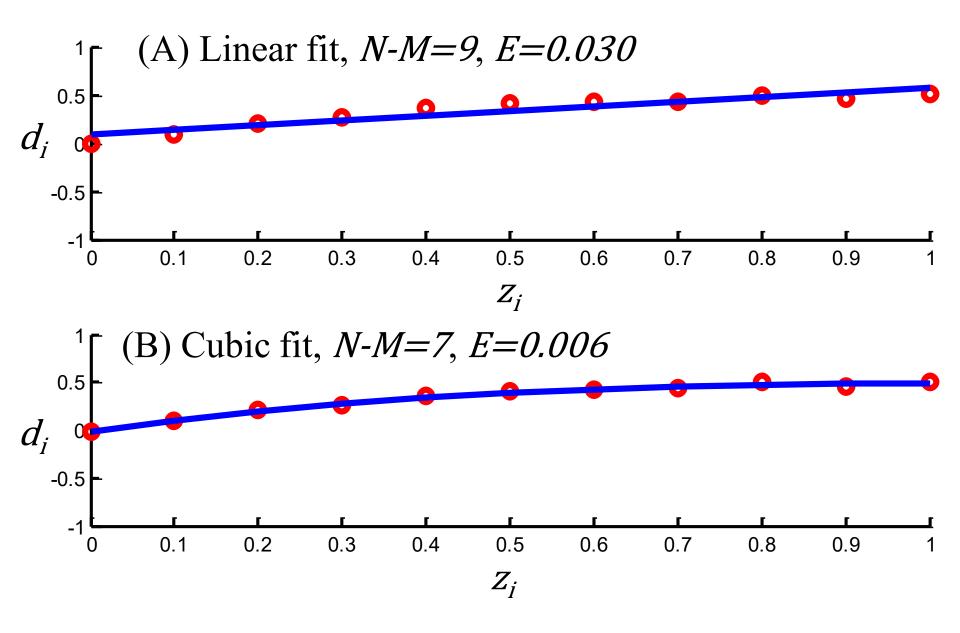
$$\langle F \rangle = \frac{\nu_B}{\nu_B - 2}$$
 $\sigma_F^2 = \frac{2\nu_B^2(\nu_A + \nu_B - 2)}{\nu_A(\nu_B - 2)^2(\nu_B - 4)}$

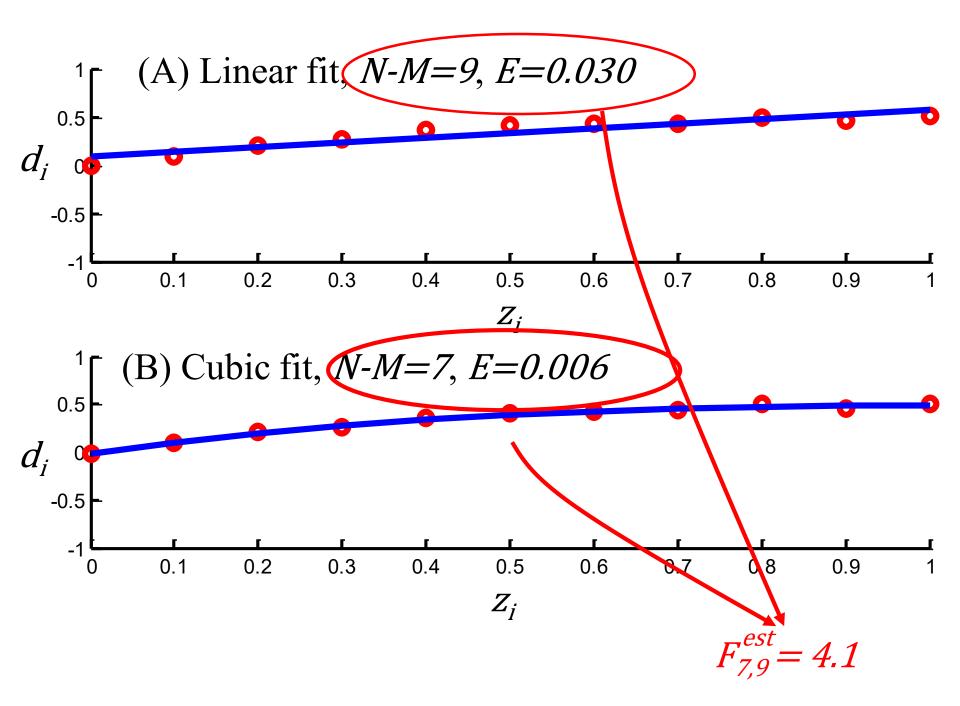
0

example

same dataset fit with

a straight line and a cubic polynomial





 $P(F < 1/F^{est} \text{ or } F > F^{est})$

probability that $F > F^{est}$

(cubic fit seems better than linear fit) by random chance alone

or $F < 1/F^{est}$ (linear fit seems better than cubic fit) by random chance alone

in MatLab

P = 1 - (fcdf(Fobs,vA,vB)-fcdf(1/Fobs,vA,vB));

answer: 6%

The Null Hypothesis

that the difference is due to random variation

cannot be rejected to 95% confidence