

Backus-Gilbert Resolution Computed from Ensembles
Bill Menke, October 31, 2018

Suppose the model parameter vector \mathbf{m} (of length M) and its covariance \mathbf{C}_m are estimated from an ensemble of solutions. Writing the eigenvalue expansion of the covariance matrix as $\mathbf{C}_m = \mathbf{V}\mathbf{\Lambda}\mathbf{V}^T$, we consider the linear transformation $\mathbf{d} = \mathbf{G}\mathbf{m}$ with $\mathbf{G} \equiv \mathbf{V}^T$. The covariance matrix \mathbf{C}_d of \mathbf{d} is:

$$\mathbf{C}_d = \mathbf{G}\mathbf{C}_m\mathbf{G}^T = \mathbf{V}^T\mathbf{V}\mathbf{\Lambda}\mathbf{V}^T\mathbf{V}^T = \mathbf{\Lambda}$$

That is, the \mathbf{d} 's are uncorrelated with covariance $\mathbf{\Lambda}$. Now suppose that we discard all eigenvectors $\mathbf{v}^{(i)}$ for which $\Lambda_{ii} > \sigma^2$, where σ is a threshold variance, and create an eigenvalue matrix \mathbf{V}_P^T of the remaining, say P , eigenvectors. The linear combinations $\mathbf{d}_P = \mathbf{G}_P\mathbf{m}$ with $\mathbf{G}_P \equiv \mathbf{V}_P^T$ all have variance above the threshold.

Now suppose that we define an average $a = \mathbf{a}^T\mathbf{d}_P$, with the averaging vector \mathbf{a} chosen to have unit length; that is $\mathbf{a}^T\mathbf{a} = 1$. Then the variance σ_a^2 of a can be no larger than the threshold σ^2 :

$$\sigma_a^2 = \mathbf{a}^T\mathbf{C}_g\mathbf{a} = \sum_{i=1}^P a_i^2 \Lambda_{ii} \leq \max_j \Lambda_{jj} \sum_{i=1}^P a_i^2 = \sigma^2$$

The Backus-Gilbert (1967) method allows one to determine an averaging vector $\mathbf{a}^{(k)}$ associated with a linear combination of model that is “most-localized” around a given model parameter m_k . The goal is to make the resolution vector $\mathbf{r}^{(k)}$ in:

$$\mathbf{a}^{(k)} = [\mathbf{a}^{(k)}]^T \mathbf{g}_P = \{[\mathbf{a}^{(k)}]^T \mathbf{G}_P\} \mathbf{m} \equiv [\mathbf{r}^{(k)}]^T \mathbf{m}$$

as close to a “unit spike” as possible, so that $a^{(k)}$ can be interpreted as the average of just a few model parameters, centered about m_k . The solution minimizes a measure J_k of the spread of $\mathbf{r}^{(k)}$:

$$\text{minimize } J_k = \sum_i w(l, k) r_l^{(k)} r_l^{(k)} \quad \text{where } \mathbf{r}^{(k)} \equiv [\mathbf{a}^{(k)}]^T \mathbf{G}_P \quad \text{and} \quad \sum_i a_i^{(k)} = 1$$

Here $w(l, k)$ is a penalty function that quantifies the distance between model parameters m_l and m_k and that satisfied $w(l, l) = 0$ and $w(l, k) > 0$ for $l \neq k$. Backus and Gilbert (1967) showed that the solution to this problem is:

$$a_i^{(k)} = \frac{\sum_i u_i [\{\mathbf{S}^{(k)}\}^{-1}]_{il}}{\sum_i \sum_i u_i [\{\mathbf{S}^{(k)}\}^{-1}]_{ij} u_j} \quad \text{with} \quad u_j = \sum_i [\mathbf{G}_P]_{jk}$$

$$\text{and } [\mathbf{S}^{(k)}]_{ij} = \sum_l w(l, k) [\mathbf{G}_P]_{il} [\mathbf{G}_P]_{jl}$$

The function $J_k(\sigma)$ defines how spread of resolution and variance trade off. Note that the Backus-Gilbert constraint on the size of the elements of $\mathbf{a}^{(k)}$ are different than the one discussed in the context of variance, so the upper limit on variance is, at best, only approximate.