

Backus-Gilbert Resolution Computed from Ensembles  
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Suppose the model parameter vector  $\mathbf{m}$  (of length  $M$ ) and its covariance  $\mathbf{C}_m$  are estimated from an ensemble of solutions. Writing the eigenvalue expansion of the covariance matrix as  $\mathbf{C}_m = \mathbf{V}\Lambda\mathbf{V}^T$ , we consider the linear transformation  $\mathbf{d} = \mathbf{G}\mathbf{m}$  with  $\mathbf{G} \equiv \mathbf{V}^T$ . The covariance matrix  $\mathbf{C}_d$  of  $\mathbf{d}$  is:

$$\mathbf{C}_d = \mathbf{G}\mathbf{C}_m\mathbf{G}^T = \mathbf{V}^T\mathbf{V}\Lambda\mathbf{V}^T\mathbf{V}^T = \Lambda$$

That is, the  $\mathbf{d}$ 's are uncorrelated with covariance  $\Lambda$ . Now suppose that we discard all eigenvectors  $\mathbf{v}^{(i)}$  for which  $\Lambda_{ii} > \sigma^2$ , where  $\sigma$  is a threshold variance, and create an eigenvalue matrix  $\mathbf{V}_P^T$  of the remaining, say  $P$ , eigenvectors. The linear combinations  $\mathbf{d}_P = \mathbf{G}_P\mathbf{m}$  with  $\mathbf{G}_P \equiv \mathbf{V}_P^T$  all have variance above the threshold.

Now suppose that we define an average  $a = \mathbf{a}^T\mathbf{d}_P$ , with the averaging vector  $\mathbf{a}$  chosen to have unit length; that is  $\mathbf{a}^T\mathbf{a} = 1$ . Then the variance  $\sigma_a^2$  of  $a$  can be no larger than the threshold  $\sigma^2$ :

$$\sigma_a^2 = \mathbf{a}^T \mathbf{C}_g \mathbf{a} = \sum_{i=1}^P a_i^2 \Lambda_{ii} \leq \max_j \Lambda_{jj} \sum_{i=1}^P a_i^2 = \sigma^2$$

The Backus-Gilbert (1967) method allows one to determine an averaging vector  $\mathbf{a}^{(k)}$  associated with a linear combination of model that is “most-localized” around a given model parameter  $m_k$ . The goal is to make the resolution vector  $\mathbf{r}^{(k)}$  in:

$$a^{(k)} = [\mathbf{a}^{(k)}]^T \mathbf{g}_P = \{[\mathbf{a}^{(k)}]^T \mathbf{G}_P\} \mathbf{m} \equiv [\mathbf{r}^{(k)}]^T \mathbf{m}$$

as close to a “unit spike” as possible, so that  $a^{(k)}$  can be interpreted as the average of just a few model parameters, centered about  $m_k$ . The solution minimizes a measure  $J_k$  of the spread of  $\mathbf{r}^{(k)}$ :

$$\text{minimize } J_k = \sum_l w(l, k) r_l^{(k)} r_l^{(k)} \quad \text{where } \mathbf{r}^{(k)} \equiv [\mathbf{a}^{(k)}]^T \mathbf{G}_P \quad \text{and} \quad \sum_i a_i^{(k)} = 1$$

Here  $w(l, k)$  is a penalty function that quantifies the distance between model parameters  $m_l$  and  $m_k$  and that satisfied  $w(l, l) = 0$  and  $w(l, k) > 0$  for  $l \neq k$ . Backus and Gilbert (1967) showed that the solution to this problem is:

$$a_l^{(k)} = \frac{\sum_i u_i [\{\mathbf{S}^{(k)}\}^{-1}]_{il}}{\sum_i \sum_l u_i [\{\mathbf{S}^{(k)}\}^{-1}]_{ij} u_j} \quad \text{with} \quad u_j = \sum_i [\mathbf{G}_P]_{jk}$$

$$\text{and} \quad [\mathbf{S}^{(k)}]_{ij} = \sum_l w(l, k) [\mathbf{G}_P]_{il} [\mathbf{G}_P]_{jl}$$

The function  $J_k(\sigma)$  defines how spread of resolution and variance trade off. Note that the Backus-Gilbert constraint on the size of the elements of  $\mathbf{a}^{(k)}$  are different than the one discussed in the context of variance, so the upper limit on variance is, at best, only approximate.