

EESC 9945

Geodesy with the Global Positioning
System

Classes 11: *Stochastic Filters II*

Review: Kalman Filter equations

Prediction

$$\hat{x}_{k|k-1} = S_k \hat{x}_{k-1|k-1}$$

$$\Lambda_{k|k-1} = S_k \Lambda_{k-1|k-1} S_k^T + R_k Q_k R_k^T$$

Gain

$$K_k = \Lambda_{k|k-1} A_k^T (A_k \Lambda_{k|k-1} A_k^T + G_k)^{-1}$$

Update

$$\hat{x}_{k|k} = \hat{x}_{k|k-1} + K_k \Delta y_k$$

$$\Lambda_{k|k} = (I - K_k A_k) \Lambda_{k|k-1}$$

Example: Random walk

- Let us take the simple 1-D example

$$z_{k+1} = z_k + \xi_k$$

ξ_k a zero-mean white-noise process with variance σ_ξ^2

- z_k is a **random-walk** process

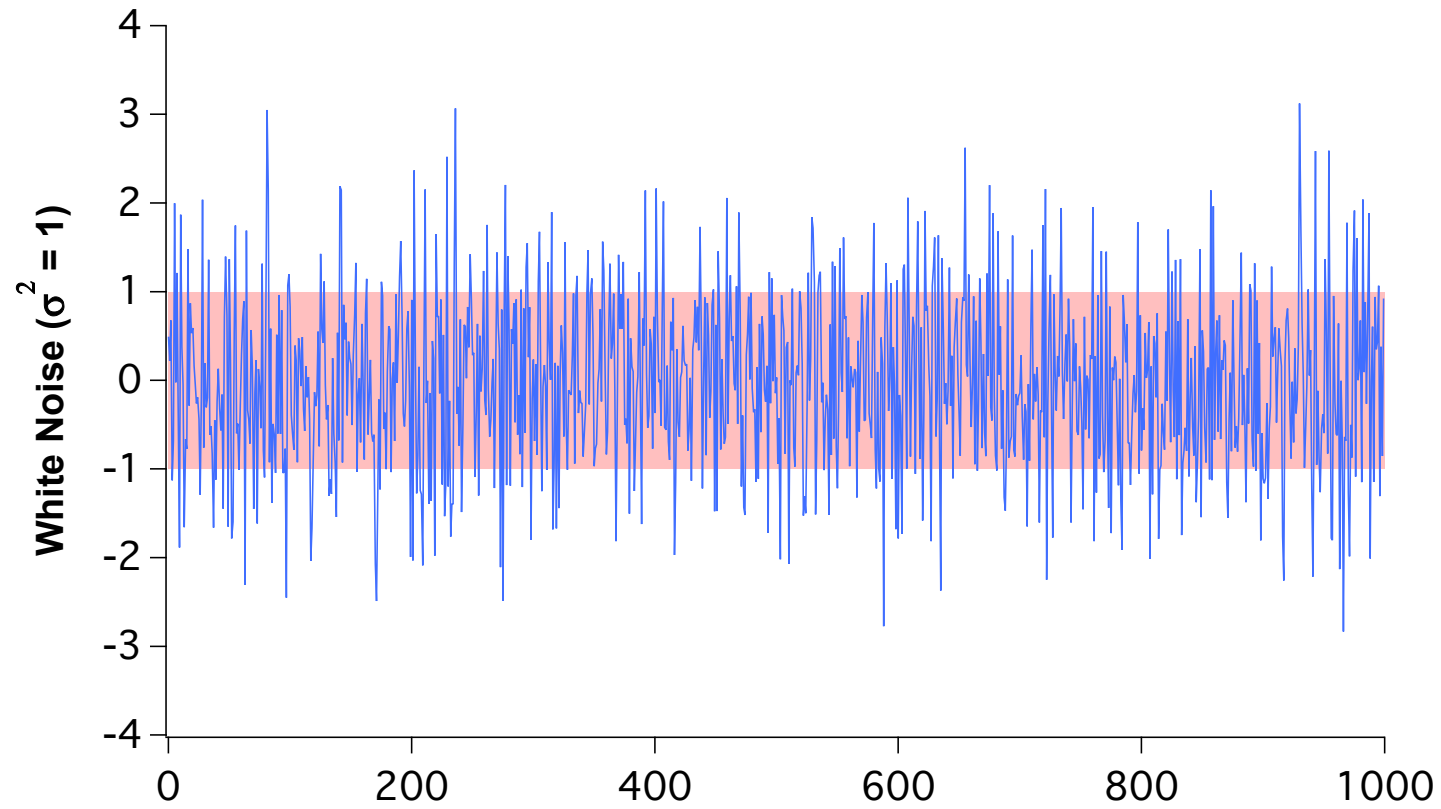
- Let us also assume an observation

$$y_k = z_k + \epsilon_k$$

with ϵ_k our observation error: $\langle \epsilon_k^2 \rangle = \sigma_y^2$

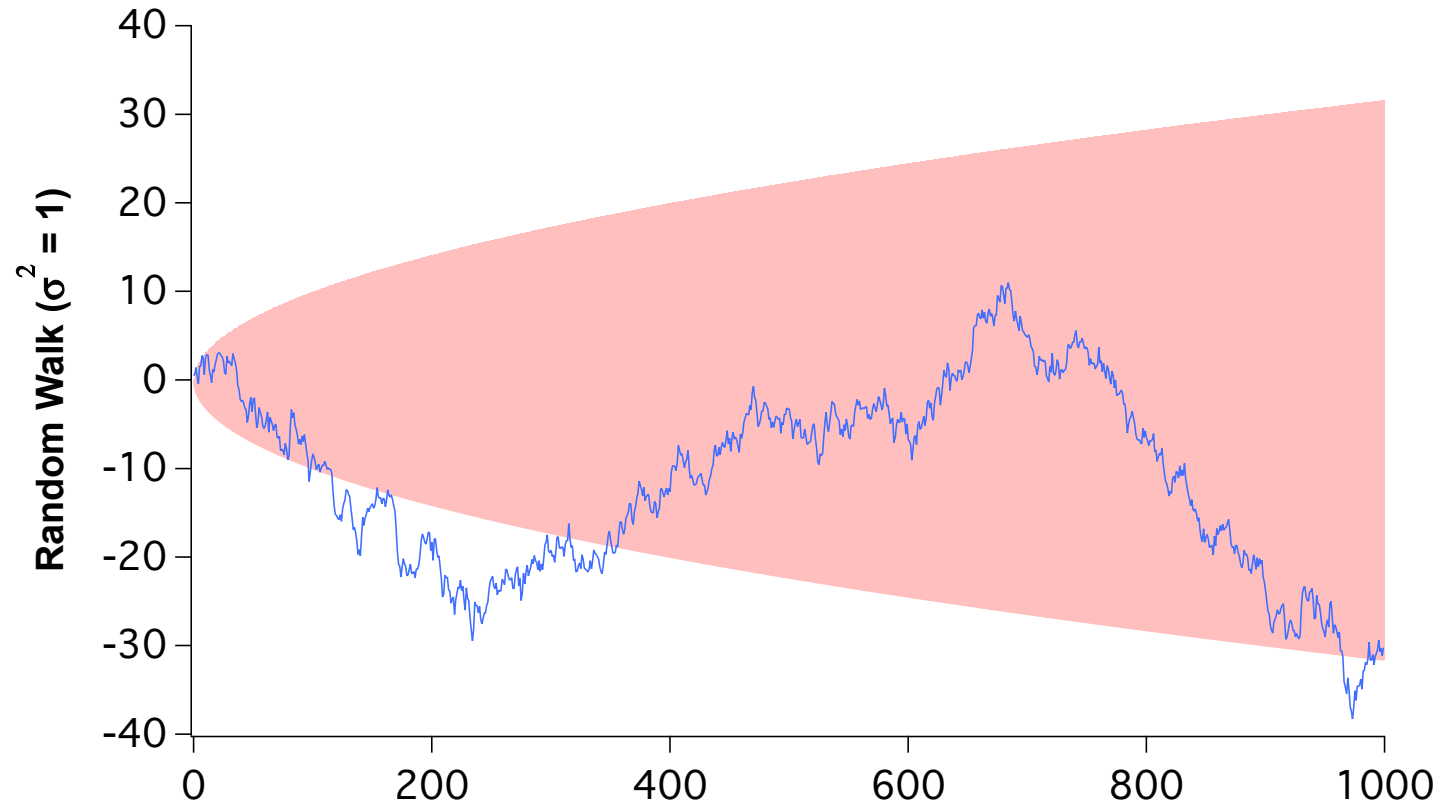
- Then $S_k = R_k = A_k = I_{1 \times 1} = 1$

White Noise Process



± 1 standard deviation region shown

Random Walk Process



± 1 standard deviation region shown

Random walk

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Random walk

- The prediction is

$$\hat{z}_{k|k-1} = \hat{z}_{k-1|k-1}$$

$$\sigma_{k|k-1}^2 = \sigma_{k-1|k-1}^2 + \sigma_{\xi}^2$$

- The Kalman gain is

$$K_k = \frac{\sigma_{k-1|k-1}^2 + \sigma_{\xi}^2}{\sigma_{k-1|k-1}^2 + \sigma_{\xi}^2 + \sigma_y^2}$$

Random walk

- The update is

$$\hat{z}_{k|k} = \hat{z}_{k-1|k-1} + \left(\frac{\sigma_{k-1|k-1}^2 + \sigma_{\xi}^2}{\sigma_{k-1|k-1}^2 + \sigma_{\xi}^2 + \sigma_y^2} \right) [y_k - \hat{z}_{k-1|k-1}]$$

$$\sigma_{k|k}^2 = \left(\frac{\sigma_y^2}{\sigma_{k-1|k-1}^2 + \sigma_{\xi}^2 + \sigma_y^2} \right) \sigma_{k-1|k-1}^2$$

Random walk

$$\hat{z}_{k|k} = \hat{z}_{k-1|k-1} + \left(\frac{\sigma_{k-1|k-1}^2 + \sigma_{\xi}^2}{\sigma_{k-1|k-1}^2 + \sigma_{\xi}^2 + \sigma_y^2} \right) [y_k - \hat{z}_{k-1|k-1}]$$

$$\sigma_{k|k}^2 = \left(\frac{\sigma_y^2}{\sigma_{k-1|k-1}^2 + \sigma_{\xi}^2 + \sigma_y^2} \right) \sigma_{k-1|k-1}^2$$

- Suppose we have an initial “guess” of $\hat{z}_{0|0} = Z_0$ with a large uncertainty, so $\sigma_{0|0}^2 \gg \sigma_{\xi}^2, \sigma_y^2$
- Then after the first step ($k = 1$) we have

$$\hat{z}_{1|1} \simeq y_1 \quad \sigma_{1|1}^2 \simeq \sigma_y^2$$

Random walk

- After n steps we have

$$\sigma_{n|n}^2 = \left[\frac{\beta}{(1 + \beta)^n - 1} \right] \sigma_y^2, \quad \beta = \sigma_\xi^2 / \sigma_y^2$$

- Limits:

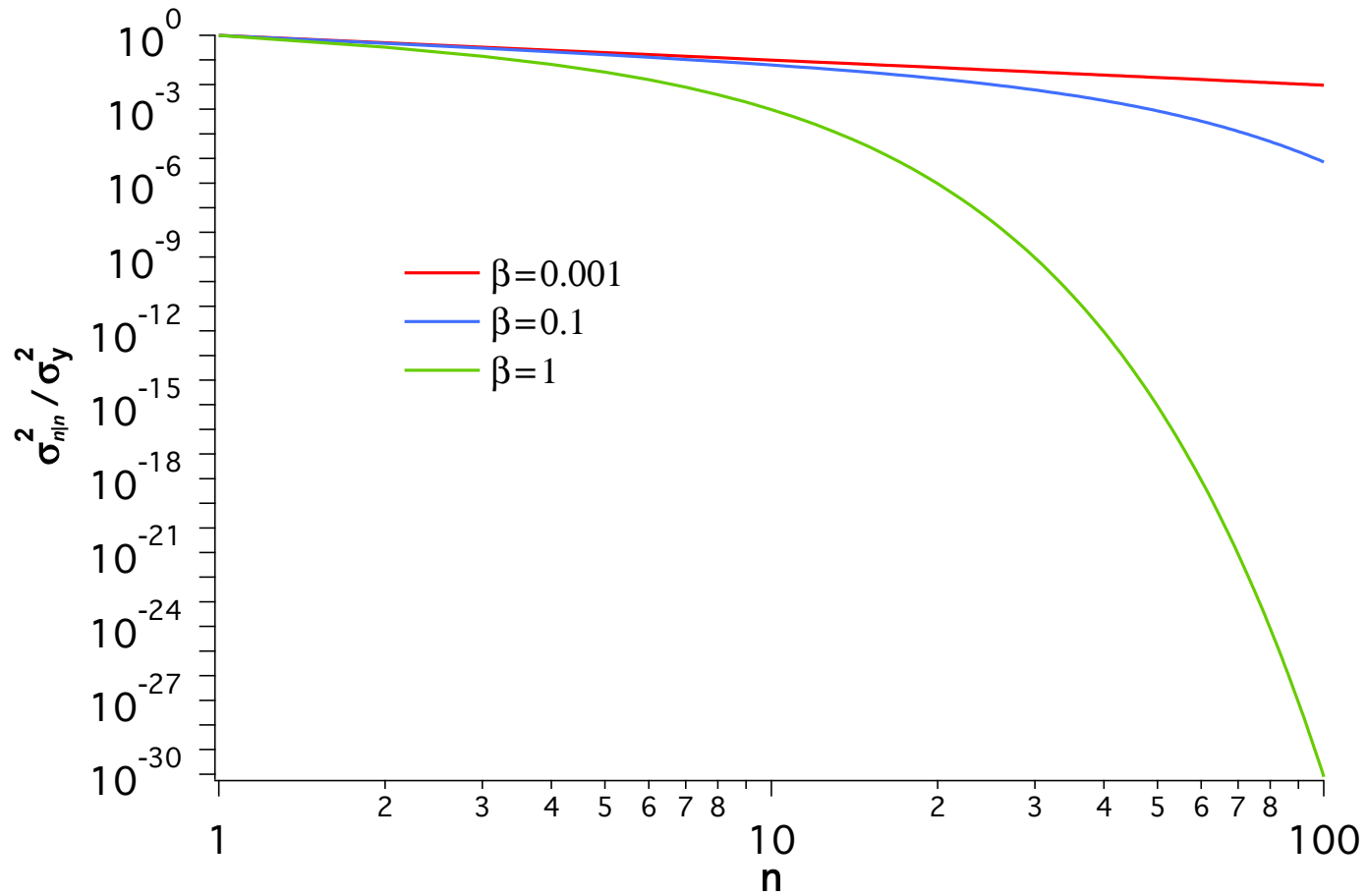
1. $\beta \rightarrow 0$ ($\sigma_\xi^2 \rightarrow 0$):

$$\sigma_{n|n}^2 / \sigma_y^2 \rightarrow \frac{1}{n}$$

2. $\beta \rightarrow \infty$ ($\sigma_y^2 \rightarrow 0$): $\sigma_{n|n}^2 \rightarrow \sigma_y^2 / \beta^{n-1}$

$$\sigma_{n|n}^2 / \sigma_y^2 \rightarrow \frac{1}{\beta^{n-1}}$$

Random-Walk Example



Random walk: Interpretation of results

- In this example, the state is a random walk ($z_{k+1} = z_k + \xi_k$), and we directly observe the state ($y_k = z_k + \epsilon_k$)
- Why does result for $\sigma_{k|k}^2$ depend on $\beta = \sigma_{\xi}^2 / \sigma_y^2$?
- Recall that the derivation of the sequential least-squares used the weighted mean of the predicted state and the state derived from the observation
- From the point of view of least-squares, each of these state estimates is a random variable, equivalent except for their standard deviations

Random walk: Interpretation of results

- A small value for β is roughly equivalent to the statement, *The state is walking around, but nearly imperceptibly compared to our ability to observe it*
- Thus, the filter acts to mostly keep the state fixed in time, and to ascribe any variability to observational error
- Hence, $\sigma_{n|n} \sim n^{-1}$ just like a constant mean value is being estimated

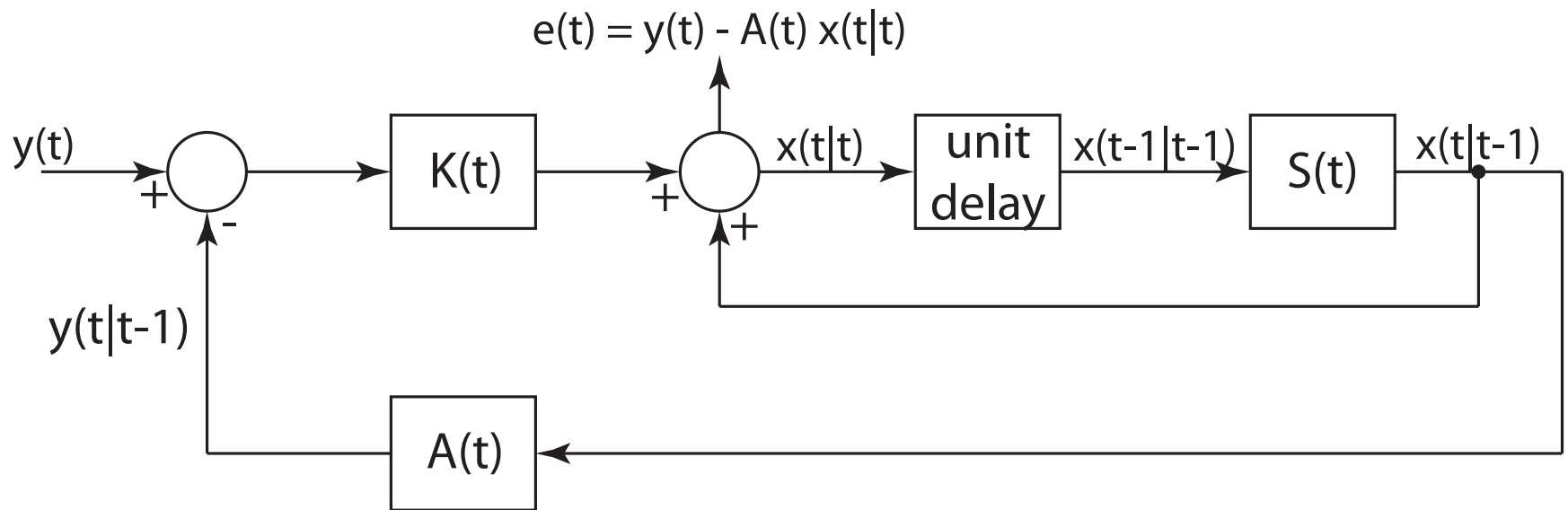
Random walk: Interpretation of results

- A large value for β is roughly equivalent to the statement, *The state is walking around, and compared to this variability we can observe it nearly perfectly*
- Thus, the filter ascribes any variability to real variability of the state itself...
- ...and zero uncertainty to error in the estimate of the state

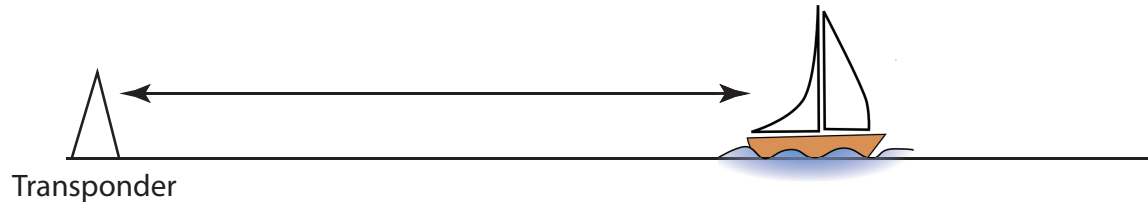
Random walk: Interpretation of results

- The preceding interpretation shows why this is called a *filter*
- The filter takes (in this example) or a random input observations), and a stochastic model, and based on the details of the model separates the signals into two components:
 - The estimate of the state vector x_k
 - The post-fit residual $\epsilon_k = y_k - A_k x_k$

Kalman Filter Schematic



1-D position with random-walk speed



- Speed varies as random walk (wind? currents?)
- Temporal variability (ξ_k zero-mean white-noise, σ_ξ^2):

$$z_k = z_{k-1} + v_{k-1} \Delta t$$

$$v_k = v_{k-1} + \xi_k$$

- Must include both position and speed in state

1-D position with random-walk speed

- Dynamic model for state:

$$x_k = \begin{bmatrix} z_k \\ v_k \end{bmatrix} = \begin{pmatrix} 1 & \Delta t \\ 0 & 1 \end{pmatrix} x_{k-1} + \begin{pmatrix} 0 \\ 1 \end{pmatrix} \xi_k$$

- Observe two-way range delay (obs. sigma σ_y^2):

$$y_k = \begin{pmatrix} \frac{2}{c} & 0 \end{pmatrix} x_k + \epsilon_k$$

- The partial derivative with respect to speed is zero!
- Can we estimate the speed?

1-D position with random-walk speed

- Kalman gain

$$K_k = \Lambda_{k|k-1} A_k^T \left(A_k \Lambda_{k|k-1} A_k^T + G_k \right)^{-1}$$

- Since $\left(A_k \Lambda_{k|k-1} A_k^T + G_k \right)^{-1}$ is 1×1 :

$$K_k \sim \Lambda_{k|k-1} A_k^T$$

- For $\Lambda_{k|k-1} = \begin{pmatrix} \sigma_z^2 & \sigma_{zv} \\ \sigma_{zv} & \sigma_v^2 \end{pmatrix}$:

$$K_k \sim \begin{pmatrix} \sigma_z^2 \\ \sigma_{zv} \end{pmatrix}$$

1-D position with random-walk speed

- Thus, there is information to estimate *both* parameters even though observation = distance

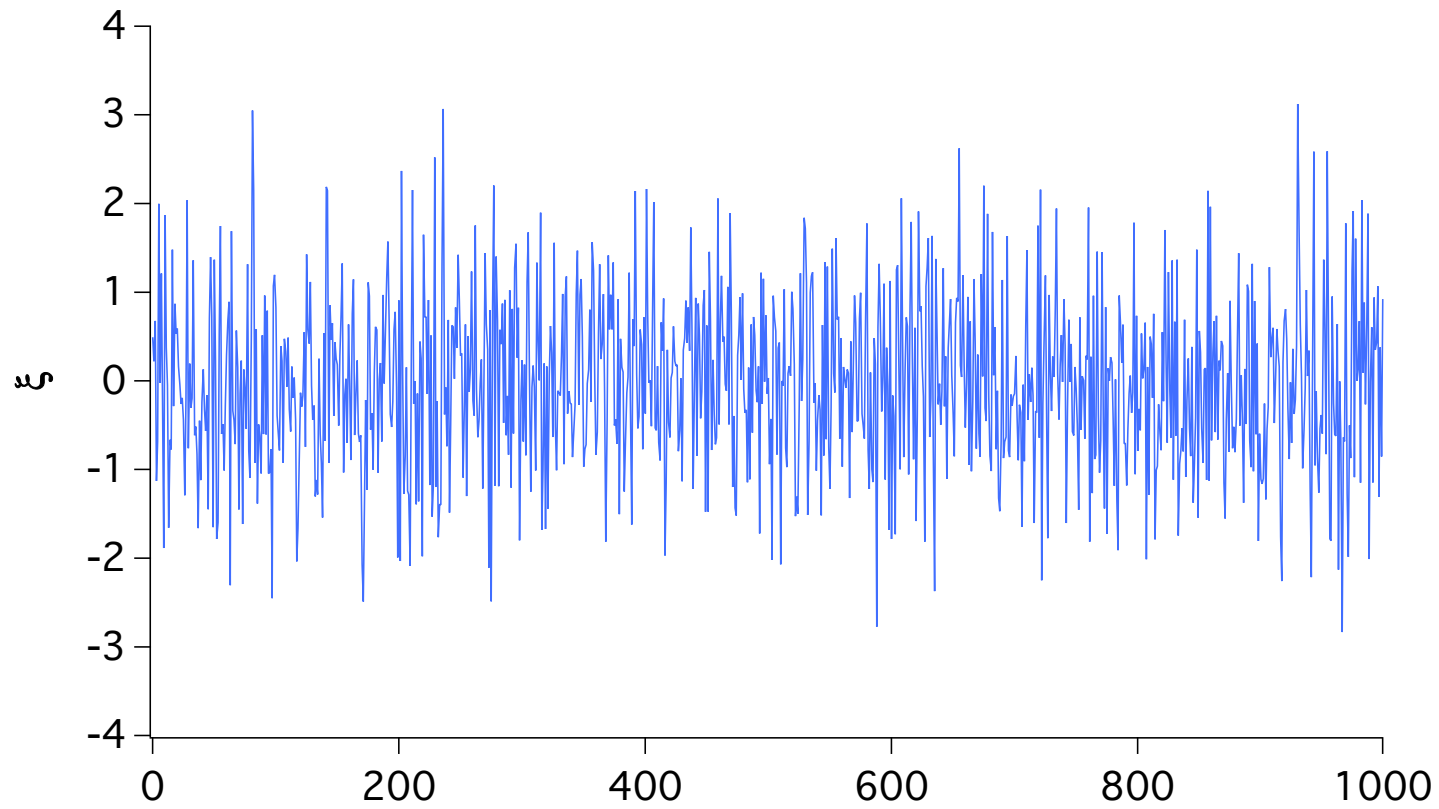
- Even on *first* epoch, if $\Lambda_{0|0} = \begin{pmatrix} \sigma_z^2 & 0 \\ 0 & \sigma_v^2 \end{pmatrix}$, then

$$\Lambda_{1|0} = \begin{pmatrix} \sigma_z^2 + (\Delta t)^2 \sigma_v^2 & \sigma_v^2 \Delta t \\ \sigma_v^2 \Delta t & \sigma_v^2 + \sigma_\xi^2 \end{pmatrix} = \begin{pmatrix} \tilde{\sigma}_z^2 & \sigma_v^2 \Delta t \\ \sigma_v^2 \Delta t & \sigma_v^2 + \sigma_\xi^2 \end{pmatrix}$$

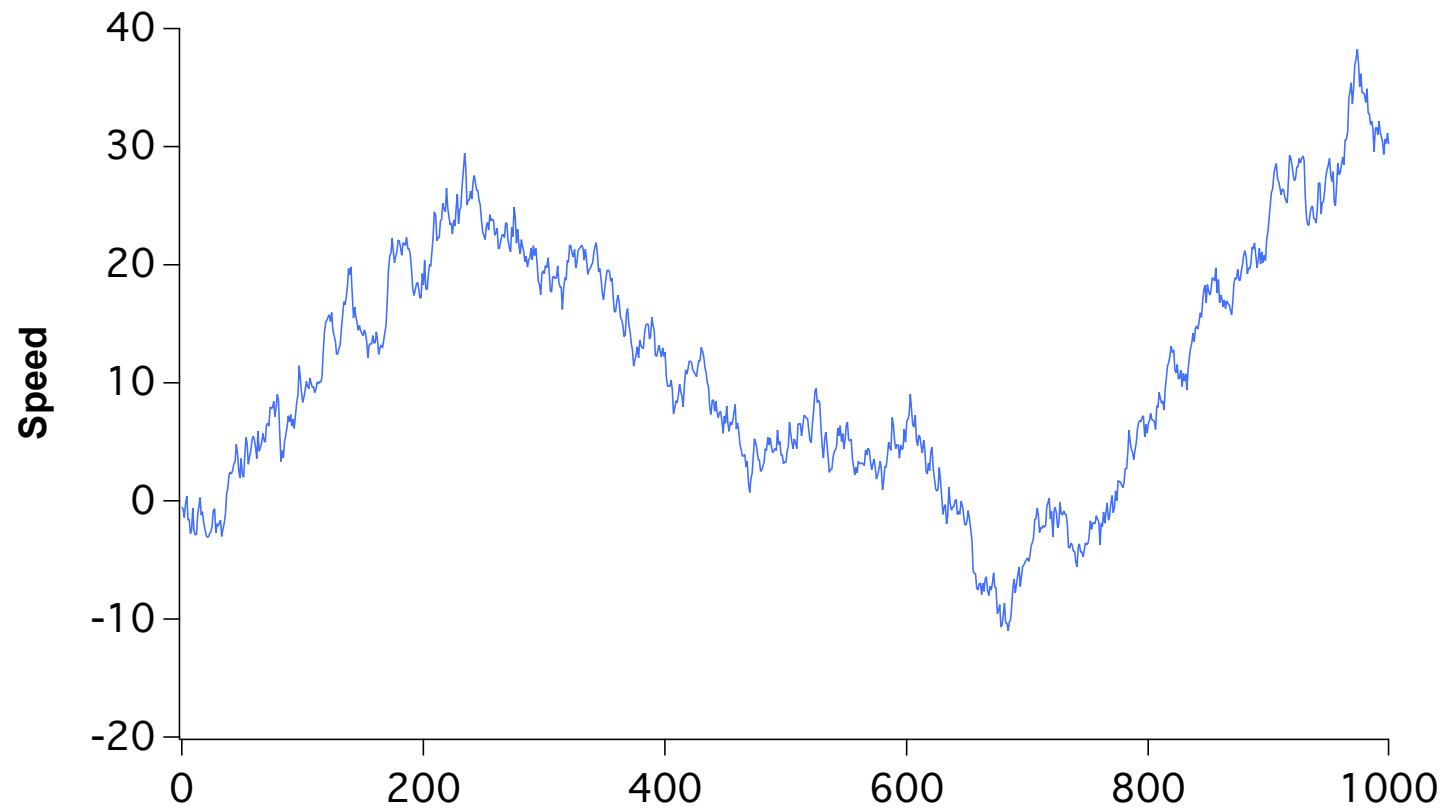
- And Kalman gain

$$K_1 = \begin{pmatrix} c \\ 2 \end{pmatrix} \left[\tilde{\sigma}_z^2 + \left(\frac{c^2}{4} \right) \sigma_v^2 \right]^{-1} \begin{pmatrix} \tilde{\sigma}_z^2 \\ \sigma_v^2 \Delta t \end{pmatrix}$$

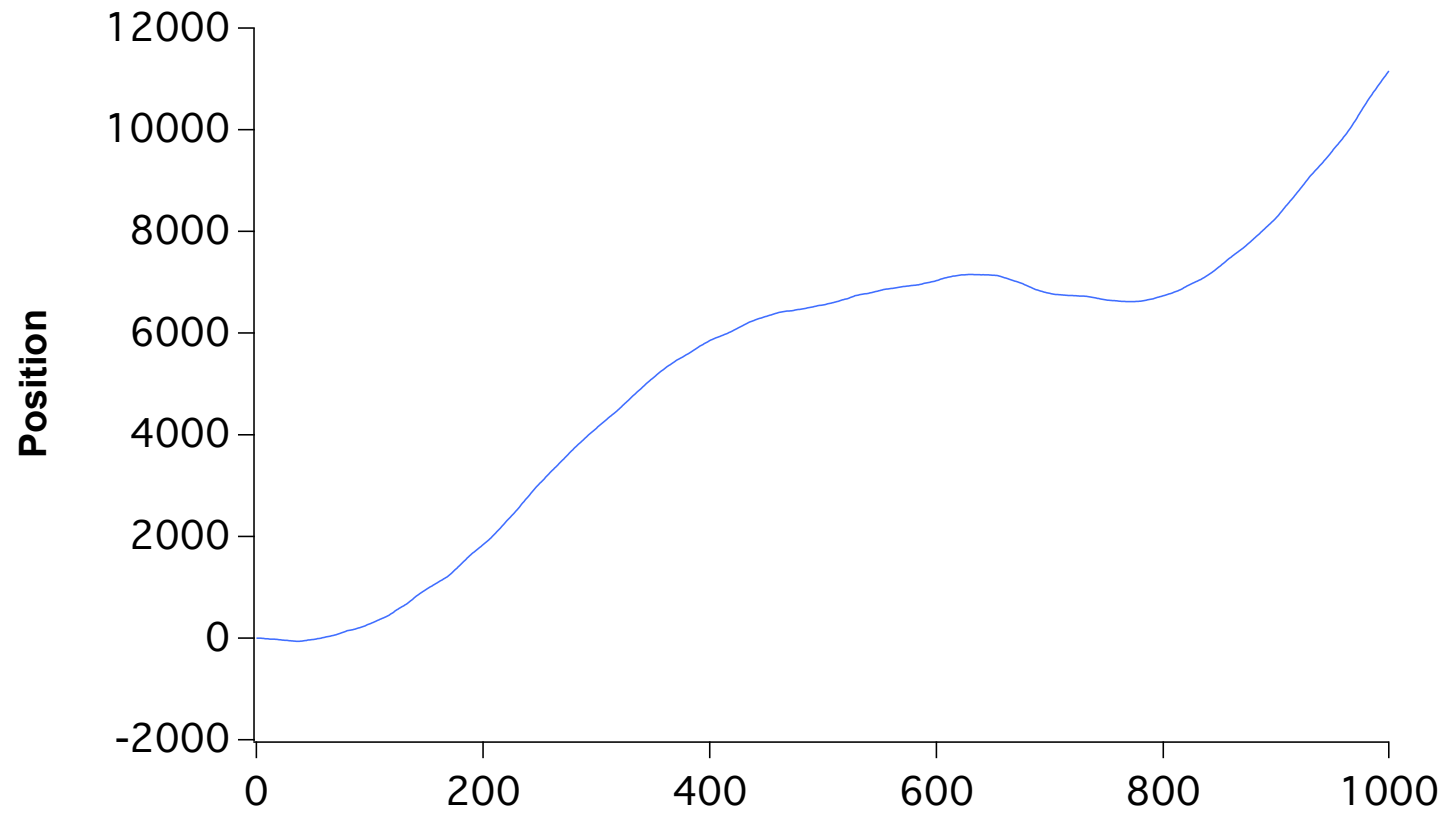
ξ (White Noise)



Speed (Random Walk)



Position (Integrated Random Walk)



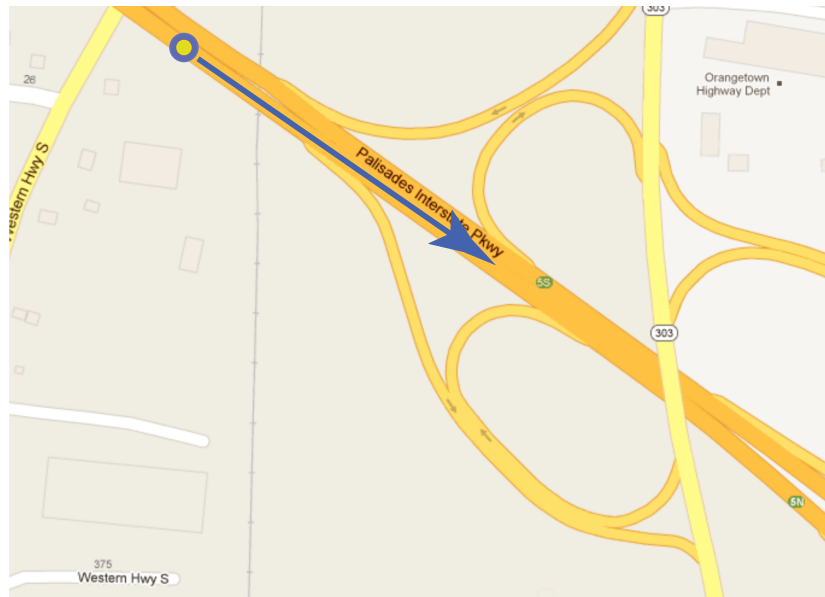
Application to GPS positioning

- Stochastic parametrization is used by various software packages for
 - Clock errors
 - Atmospheric (wet) zenith delay
 - Zenith-delay gradient parameters
 - Time-dependent positions
 - Earth orientation/rotation
 - Orbit parameters

Stochastic Models in GPS

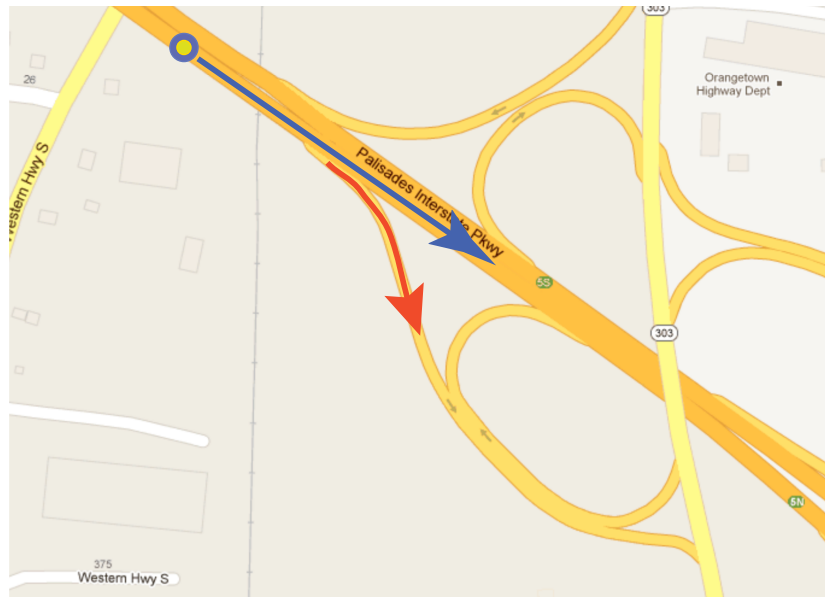
- White noise, random walk, and integrated random walk are most used
 - They are easy to implement in a Kalman filter
 - Nature of variations can be tuned to understanding of physical process
- Stochastic variances have to be “guesstimated,” and sensitivity studies done

GPS Car Navigation Example



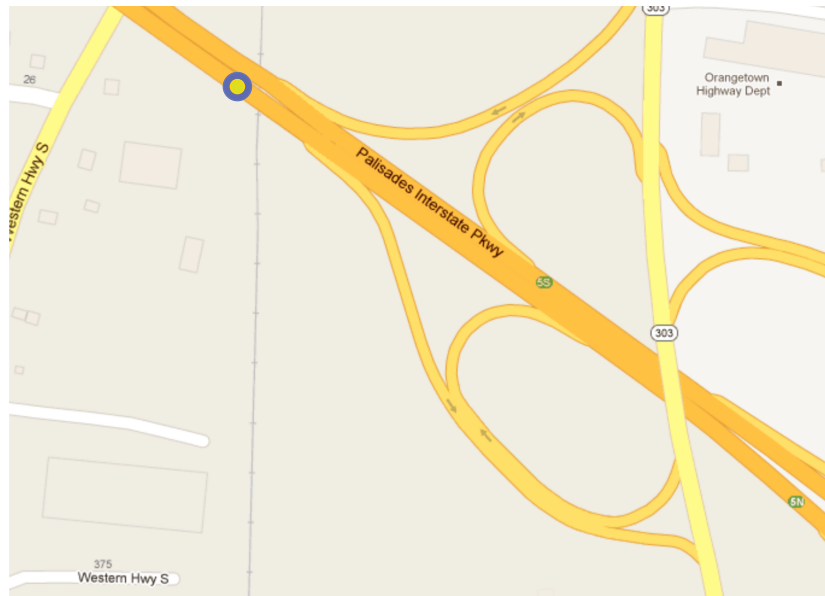
Position of your vehicle and the navigation system's route

GPS Car Navigation Example



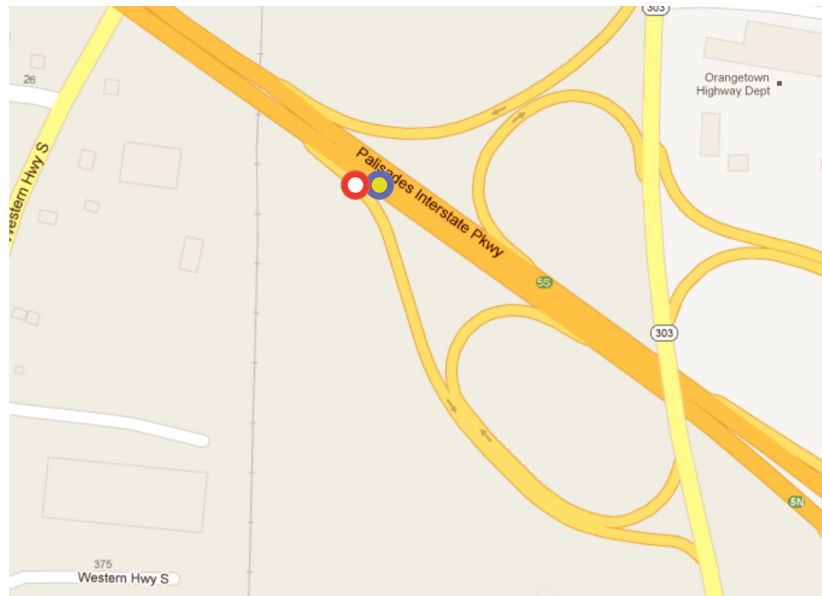
Route you want to take

GPS Car Navigation Example



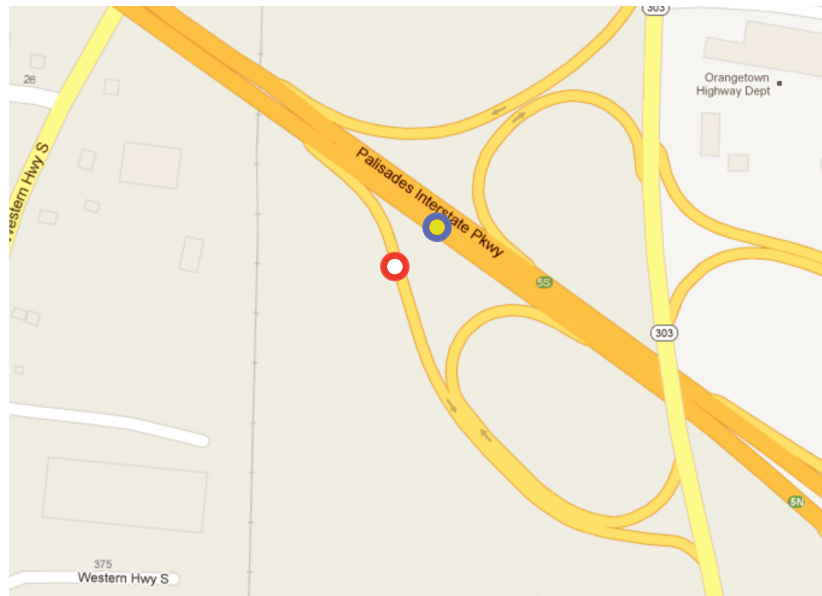
Estimated position (blue/yellow) vs. true position (red/white)

GPS Car Navigation Example



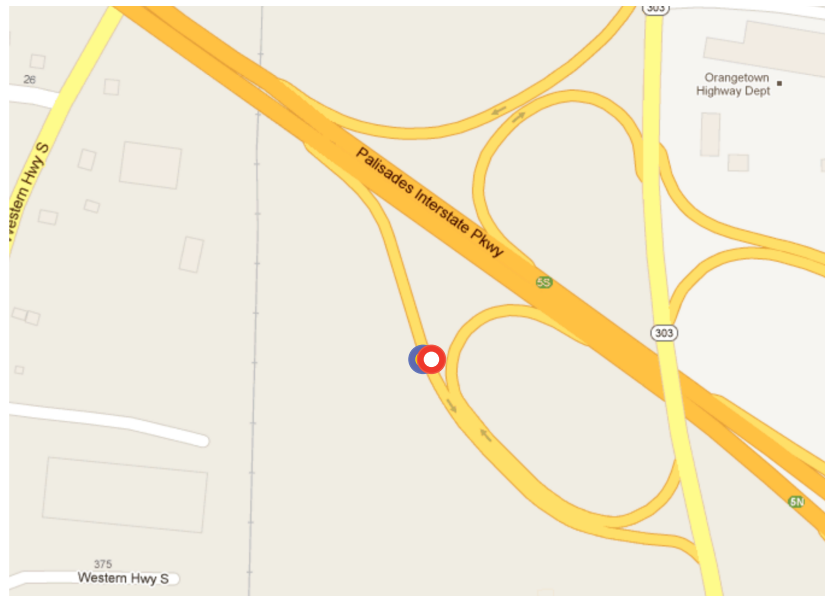
Estimated position (blue/yellow) vs. true position (red/white)

GPS Car Navigation Example



Estimated position (blue/yellow) vs. true position (red/white)

GPS Car Navigation Example



Estimated position (blue/yellow) vs. true position (red/white)

What's happening?

- Kalman filter for position and velocity (?)
- Vehicle is assumed to be on road
- Identified of road enables prediction with small σ 's
- As prefit residual pseudorange ($y_k - A_k \hat{x}_{k|k-1}$) become large, estimated position changes, but *not enough to be identified as different road*
- Prefit residuals at some point become large enough that a true position update is allowed

Combining high-accuracy GPS and accelerometer data for strong-motion displacements

- Recently, Yehuda Bock (UCSD) and colleagues have experimented with combining GPS and accelerometer estimates using a Kalman filter
- There are two ways of doing this
- One is to devise a state vector including position and acceleration, for example (for each component):

$$x_k = \begin{bmatrix} z_k \\ v_k \\ a_k \end{bmatrix}$$

Combining GPS and accelerometer data

- The transition matrix would then be

$$S_k = \begin{pmatrix} 1 & \Delta t & \frac{1}{2}(\Delta t)^2 \\ 0 & 1 & \Delta t \\ 0 & 0 & 1 \end{pmatrix}$$

- The observation vector would include both position z (from GPS) and acceleration a (from the accelerometer)
- This approach would require understanding the statistics of the acceleration changes which would be modeled as noise

Combining GPS and accelerometer data

- The approach taken by Bock et al. is to use a modified dynamic equation

$$x_k = S_k x_{k-1} + B_k u_k + R_k \xi_k$$

- The state vector is

$$x_k = \begin{bmatrix} z_k \\ v_k \end{bmatrix}$$

- The transition matrix is

$$S_k = \begin{pmatrix} 1 & \Delta t \\ 0 & 1 \end{pmatrix}$$

Combining GPS and accelerometer data

$$x_k = S_k x_{k-1} + B_k u_k + R_k \xi_k$$

- The input u_k is the accelerometer reading, so that

$$B_k = \begin{bmatrix} \frac{1}{2} (\Delta t)^2 \\ \Delta t \end{bmatrix}$$

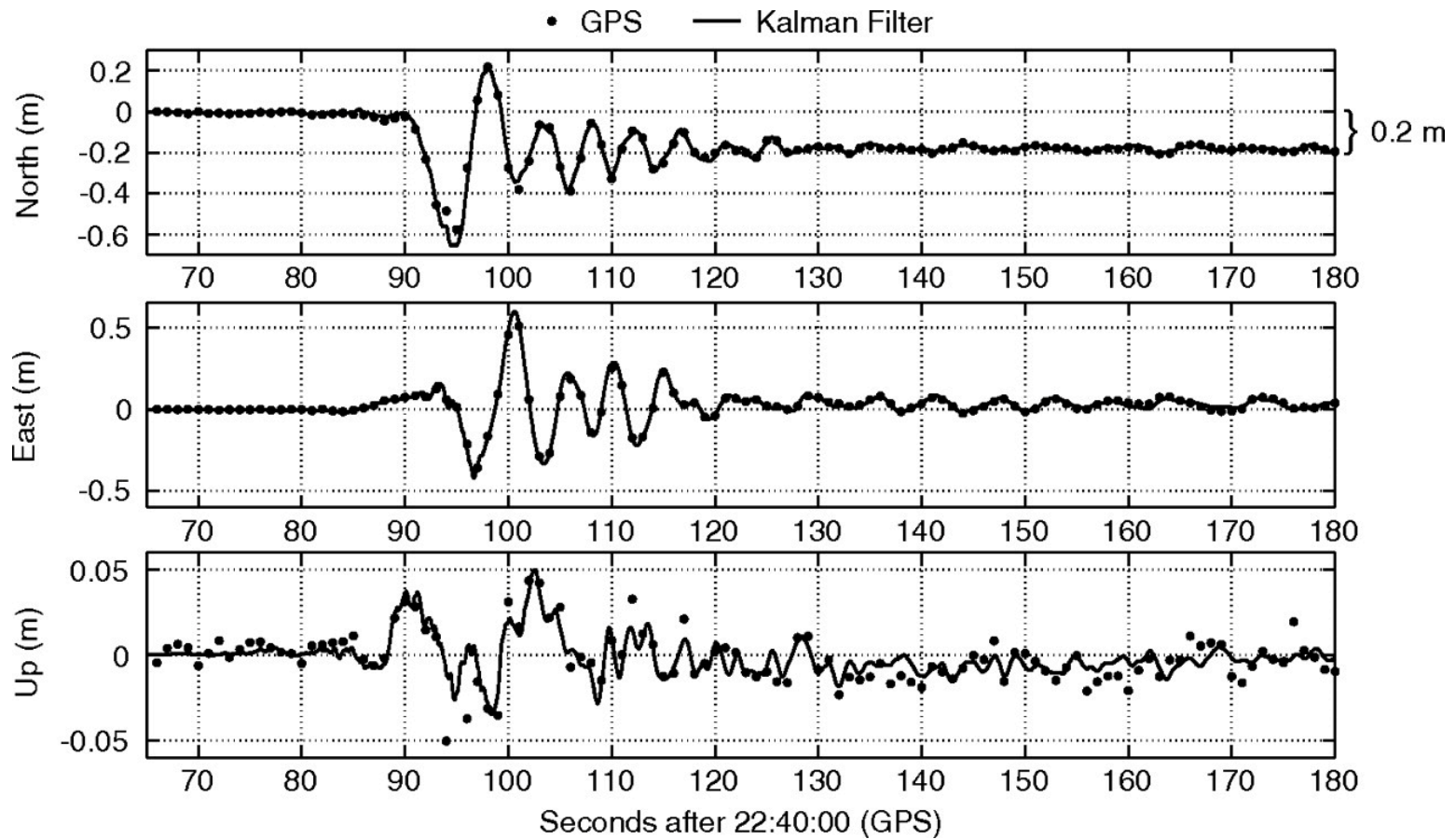
- This modified dynamic equation can be implemented by modifying the prediction equation:

$$\hat{x}_{k|k-1} = S_k x_{k-1|k-1} + B_k u_k$$

Combining GPS and accelerometer data

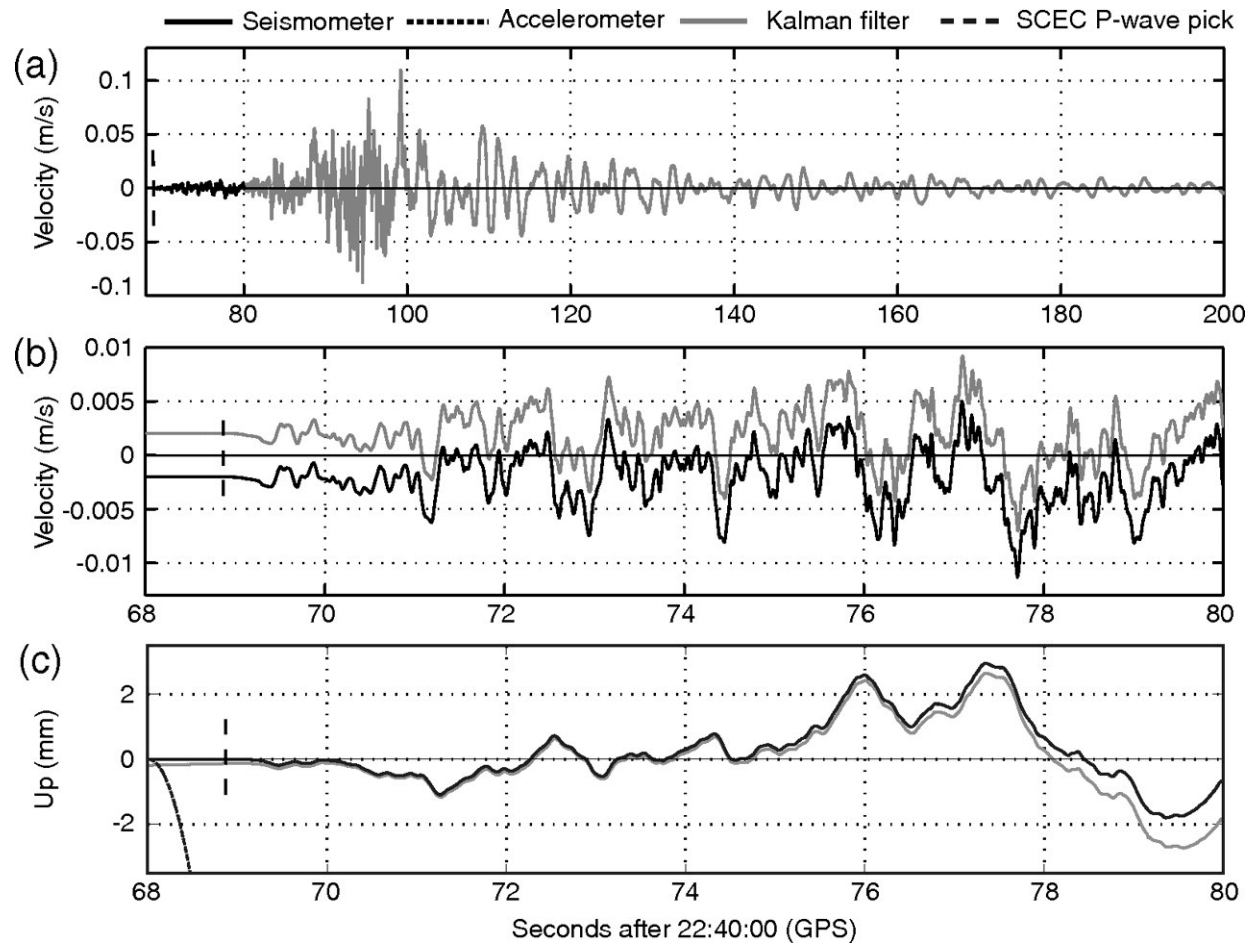
- The noise matrix Q_k has a rather complex form to reflect the sampling (with error) of a continuous quantity (acceleration)
- The GPS receiver and accelerometer have different sampling rates
 - GPS: 10 Hz
 - Accelerometer: 100 Hz

GPS vs. KF (El Mayor-Cucapah 4/4/2010)



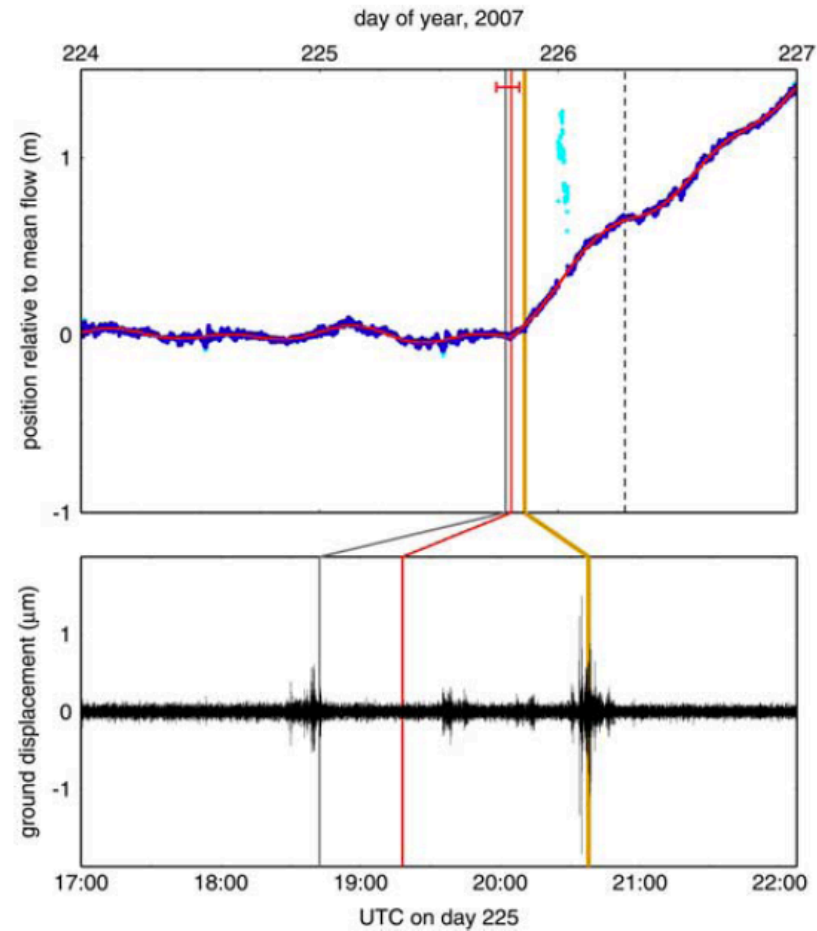
Bock et al. (2011), *BSSA*, 101, 2904–2925

Seismometer vs. KF (El Mayor-Cucapah)



Bock et al. (2011), *BSSA*, 101, 2904–2925

Position estimates of glacier GPS sites



Nettles et al. (2008), *GRL*, 35, L24503